

Hidden Markov Models for Speech Recognition

Berlin Chen 2004

References:

1. Rabiner and Juang. *Fundamentals of Speech Recognition*. Chapter 6
2. Huang et. al. *Spoken Language Processing*. Chapters 4, 8
3. Vaseghi. *Advanced Digital Signal Processing and Noise Reduction*. Chapter 5
4. Rabiner. *A Tutorial on Hidden Markov Models and Selected Applications in Speech Recognition*. Proceedings of the IEEE, vol. 77, No. 2, February 1989

Introduction

- Hidden Markov Model (HMM)

History

- Published in papers of Baum in late 1960s and early 1970s
- Introduced to speech processing by Baker (CMU) and Jelinek (IBM) in the 1970s

Assumption

- Speech signal can be characterized as a parametric random process
- Parameters can be estimated in a precise, well-defined manner

Three fundamental problems

- Evaluation of probability (likelihood) of a sequence of observations given a specific HMM
- Determination of a best sequence of model states
- Adjustment of model parameters so as to best account for observed signals

Observable Markov Model

- Observable Markov Model (Markov Chain)
 - First-order Markov chain of N states is a triple (S, A, π)
 - S is a set of N states
 - A is the $N \times N$ matrix of transition probabilities between states
 $P(s_t=j|s_{t-1}=i, s_{t-2}=k, \dots) = P(s_t=j|s_{t-1}=i) = A_{ij}$ First-order and time-invariant assumptions
 - π is the vector of initial state probability
 $\pi_j = P(s_1=j)$
 - The output of the process is the set of states at each instant of time, when each state corresponds to an observable event
 - The output in any given state is not random (**deterministic!**)
 - Too simple to describe the speech signal characteristics

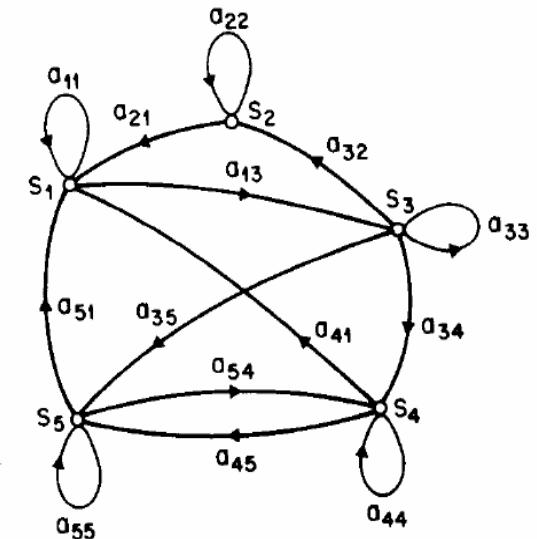


Fig. 1. A Markov chain with 5 states (labeled S₁ to S₅) with selected state transitions.

Observable Markov Model (cont.)

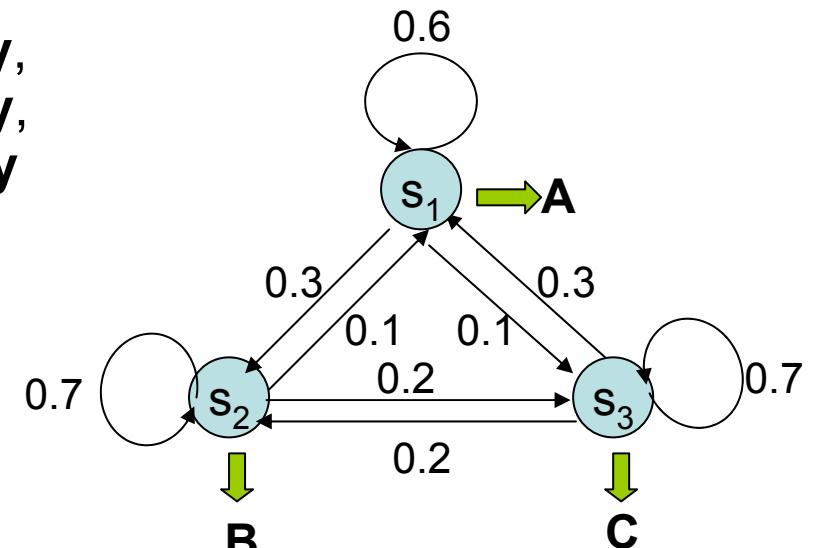
- Example 1: A 3-state Markov Chain λ

State 1 generates symbol A **only**,
State 2 generates symbol B **only**,
and State 3 generates symbol C **only**

$$A = \begin{bmatrix} 0.6 & 0.3 & 0.1 \\ 0.1 & 0.7 & 0.2 \\ 0.3 & 0.2 & 0.5 \end{bmatrix}$$

$$\pi = [0.4 \quad 0.5 \quad 0.1]$$

- Given a sequence of observed symbols $O=\{CABBCABC\}$, the **only one** corresponding state sequence is $\{S_3S_1S_2S_2S_3S_1S_2S_3\}$, and the corresponding probability is



$$P(O|\lambda)$$

$$\begin{aligned} &= P(S_3)P(S_1|S_3)P(S_2|S_1)P(S_2|S_2)P(S_3|S_2)P(S_1|S_3)P(S_2|S_1)P(S_3|S_2) \\ &= 0.1 \times 0.3 \times 0.3 \times 0.7 \times 0.2 \times 0.3 \times 0.3 \times 0.2 = 0.00002268 \end{aligned}$$

Observable Markov Model (cont.)

- Example 2: A three-state Markov chain for the *Dow Jones Industrial average*

state 1 – *up* (in comparison to the index of previous day)

state 2 – *down* (in comparison to the index of previous day)

state 3 – *unchanged* (in comparison to the index of previous day)

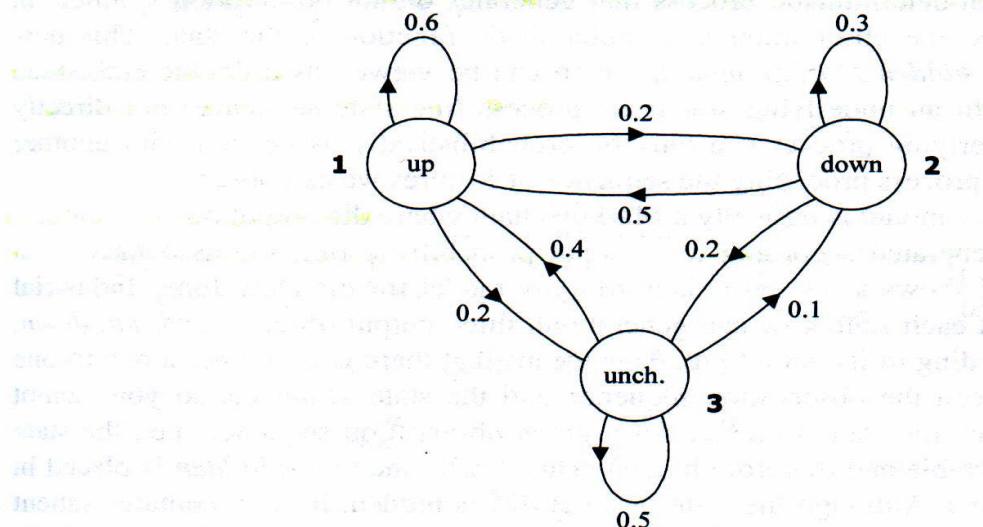


Figure 8.1 A Markov chain for the Dow Jones Industrial average. Three states represent *up*, *down*, and *unchanged*, respectively.

The parameter for this Dow Jones Markov chain may include a state-transition probability matrix

$$A = \{a_{ij}\} = \begin{bmatrix} 0.6 & 0.2 & 0.2 \\ 0.5 & 0.3 & 0.2 \\ 0.4 & 0.1 & 0.5 \end{bmatrix} \quad \pi = (\pi_i)^T = \begin{bmatrix} 0.5 \\ 0.2 \\ 0.3 \end{bmatrix}$$

and an initial state probability matrix

The probability of 5 consecutive *up* days

$$\begin{aligned} P(5 \text{ consecutive } \textit{up} \text{ days}) &= P(1,1,1,1,1) \\ &= \pi_1 a_{11} a_{11} a_{11} a_{11} a_{11} = 0.5 \times (0.6)^4 = 0.0648 \end{aligned}$$

Observable Markov Model (cont.)

- Example 3: Given a Markov model, what is the mean occupancy duration of each state i

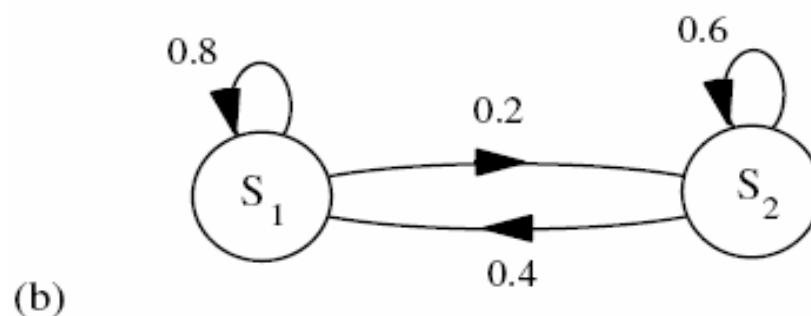
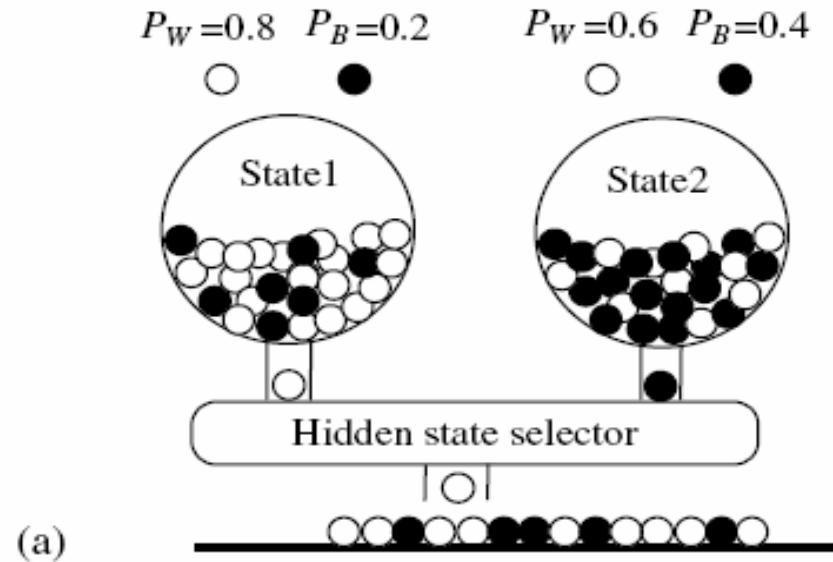
$p_i(d) = \text{prob. density function of duration } d \text{ in state } i$

$$= (a_{ii})^{d-1} (1 - a_{ii})$$

Expected number of duration in a state

$$\begin{aligned}\bar{d}_i &= \sum_{d=1}^{\infty} d p_i(d) = \sum_{d=1}^{\infty} d (a_{ii})^{d-1} (1 - a_{ii}) = (1 - a_{ii}) \frac{\partial}{\partial a_{ii}} \sum_{d=1}^{\infty} (a_{ii})^d \\ &= (1 - a_{ii}) \frac{\partial}{\partial a_{ii}} \frac{1}{1 - a_{ii}} = \frac{1}{1 - a_{ii}}\end{aligned}$$

Hidden Markov Model



(a) Illustration of a two-layered random process. (b) An HMM model of the process in (a).

Hidden Markov Model (cont.)

- HMM, an extended version of Observable Markov Model
 - The observation is turned to be a **probabilistic function (discrete or continuous) of a state** instead of an one-to-one correspondence of a state
 - The model is a **doubly embedded** stochastic process with an underlying stochastic process that is not directly observable (hidden)
 - What is hidden? **The State Sequence!**
According to the observation sequence, we are not sure which state sequence generates it!
- Elements of an HMM (the **State-Output HMM**) $\lambda=\{\mathbf{S}, \mathbf{A}, \mathbf{B}, \pi\}$
 - \mathbf{S} is a set of N states
 - \mathbf{A} is the $N \times N$ matrix of transition probabilities between states
 - \mathbf{B} is a set of N probability functions, each describing the observation probability with respect to a state
 - π is the vector of initial state probability

Hidden Markov Model (cont.)

- Two major assumptions
 - First order (Markov) assumption
 - The state transition depends only on the origin and destination
 - Time-invariant
 - Output-independent assumption
 - All observations are dependent on the state that generated them, not on neighboring observations

Hidden Markov Model (cont.)

- Two major types of HMMs according to the observations

- **Discrete and finite observations:**

- The observations that **all** distinct states generate are finite in number

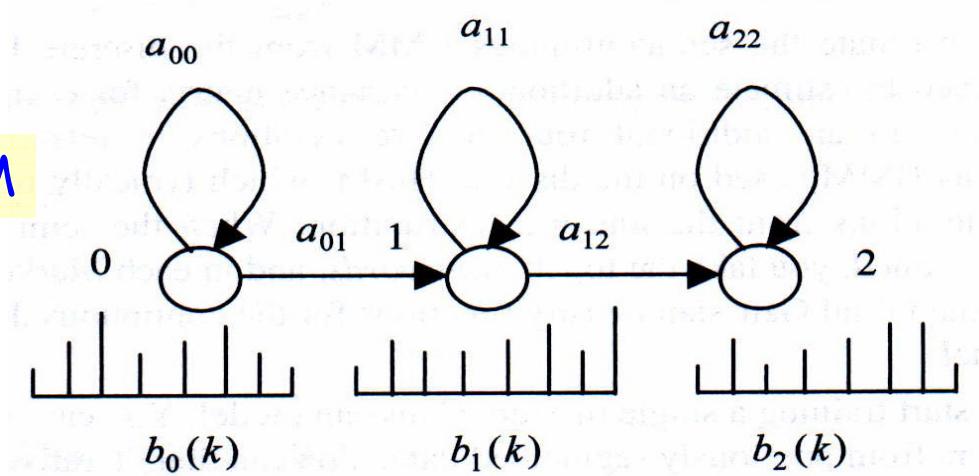
$$V = \{v_1, v_2, v_3, \dots, v_M\}, v_k \in R^L$$

- In this case, the set of observation probability distributions $B = \{b_j(v_k)\}$, is defined as $b_j(v_k) = P(o_t = v_k | s_t = j)$, $1 \leq k \leq M$, $1 \leq j \leq N$

\mathbf{o}_t : *observation at time t*, s_t : *state at time t*

\Rightarrow for state j , $b_j(v_k)$ consists of *only M probability values*

A left-to-right HMM



Hidden Markov Model (cont.)

- Two major types of HMMs according to the observations
 - Continuous and infinite observations:
 - The observations that **all** distinct states generate are infinite and continuous, that is, $V=\{v| v \in \mathbb{R}^L\}$
 - In this case, the set of observation probability distributions $B=\{b_j(v)\}$, is defined as $b_j(v)=f_{O|S}(o_t=v|s_t=j)$, $1 \leq j \leq N$
⇒ $b_j(v)$ is a **continuous probability density function (pdf)** **and is often a mixture of Multivariate Gaussian (Normal) Distributions**

$$b_j(v) = \sum_{k=1}^M w_{jk} \left(\frac{1}{(2\pi)^{L/2} |\Sigma_{jk}|^{1/2}} \exp\left(-\frac{1}{2} (v - \mu_{jk})^t \Sigma_{jk}^{-1} (v - \mu_{jk})\right) \right)$$

Covariance Matrix Mean Vector Observation Vector

Hidden Markov Model (cont.)

- Multivariate Gaussian Distributions
 - When $\mathbf{X}=(\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_L)$ is a L -dimensional random vector, the multivariate Gaussian pdf has the form:

$$f(\mathbf{X} = \mathbf{x} | \boldsymbol{\mu}, \boldsymbol{\Sigma}) = N(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{L/2} |\boldsymbol{\Sigma}|^{1/2}} \exp\left(-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^t \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right)$$

where $\boldsymbol{\mu}$ is the L -dimensional mean vector,

$\boldsymbol{\Sigma}$ is the covariance matrix, $\boldsymbol{\Sigma} = E[(\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^t] = E[\mathbf{x}\mathbf{x}^t] - \boldsymbol{\mu}\boldsymbol{\mu}^t$

and $|\boldsymbol{\Sigma}|$ is the determinant of $\boldsymbol{\Sigma}$

The $i-j^{\text{th}}$ element σ_{ij}^2 of $\boldsymbol{\Sigma}$, $\sigma_{ij}^2 = E[(x_i - \mu_i)(x_j - \mu_j)] = E[x_i x_j] - \mu_i \mu_j$

- If $\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_L$ are independent, the covariance matrix is reduced to diagonal covariance
 - The distribution as L independent scalar Gaussian distributions
 - Model complexity is reduced

Hidden Markov Model (cont.)

- Multivariate Gaussian Distributions

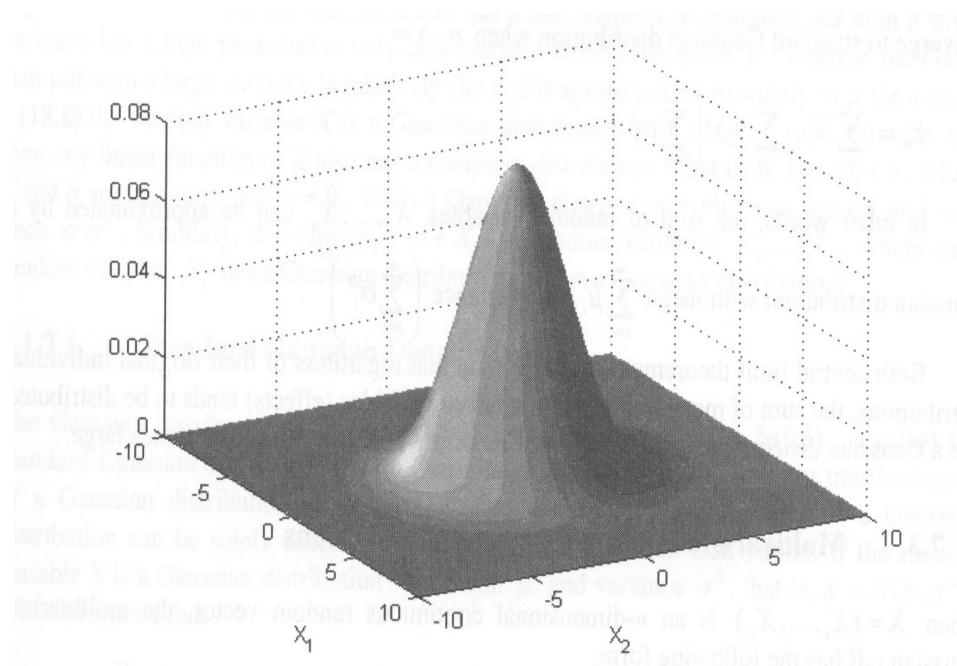


Figure 3.12 A two-dimensional multivariate Gaussian distribution with independent random variables x_1 and x_2 that have the same variance.

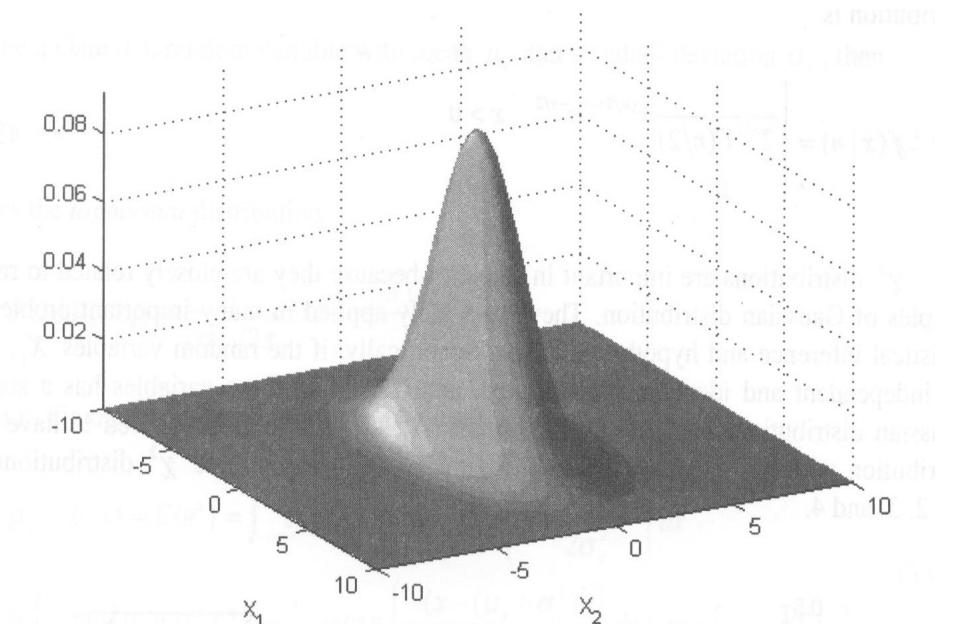
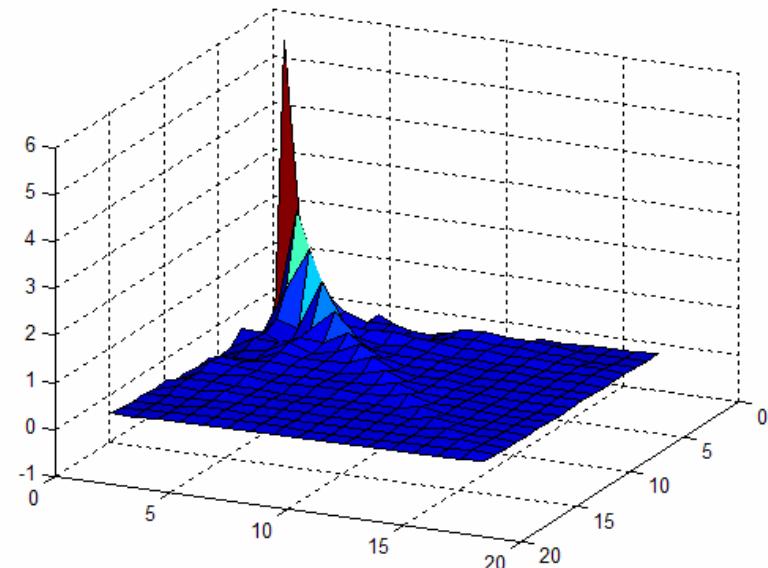
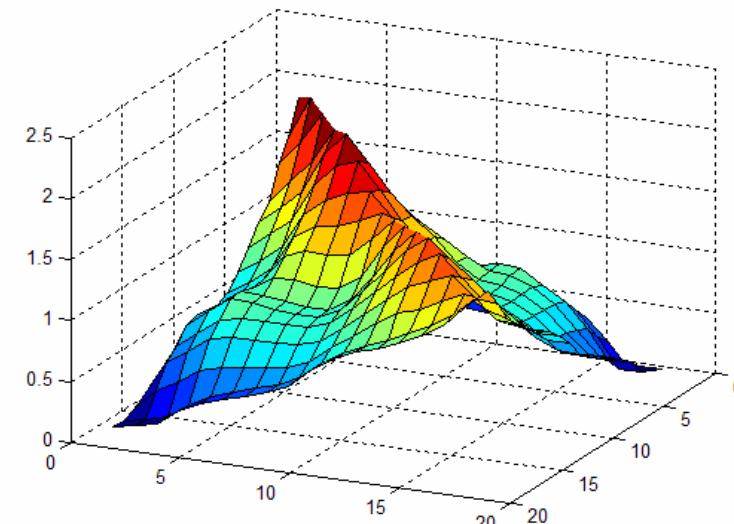


Figure 3.13 Another two-dimensional multivariate Gaussian distribution with independent random variable x_1 and x_2 which have different variances.

Hidden Markov Model (cont.)

- Covariance matrix of the correlated feature vectors (Mel Frequency filter bank outputs)
- Covariance matrix of the partially decorrelated feature vectors (MFCC cepstrum without C_0)

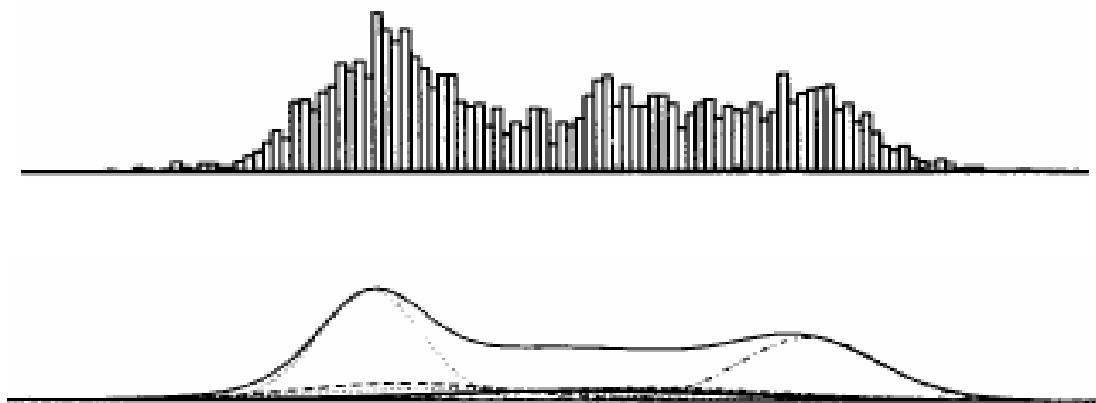


Hidden Markov Model (cont.)

- Multivariate Mixture Gaussian Distributions (cont.)
 - More complex distributions with multiple local maxima can be approximated by Gaussian (a unimodal distribution) mixture

$$f(\mathbf{x}) = \sum_{k=1}^M w_k N_k(\mathbf{x}; \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k), \quad \sum_{k=1}^M w_k = 1$$

- Gaussian mixtures with enough mixture components can approximate any distribution



Hidden Markov Model (cont.)

- Example 4: a 3-state discrete HMM λ

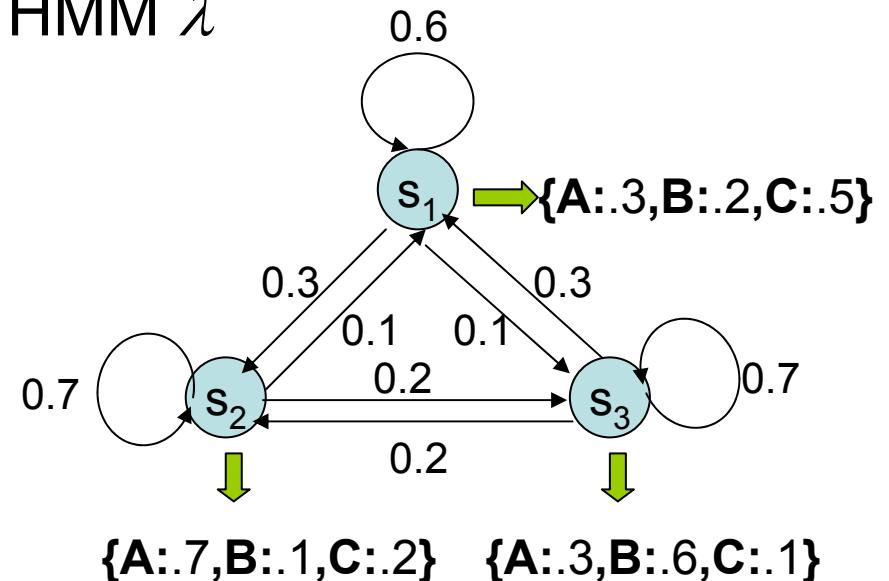
$$A = \begin{bmatrix} 0.6 & 0.3 & 0.1 \\ 0.1 & 0.7 & 0.2 \\ 0.3 & 0.2 & 0.5 \end{bmatrix}$$

$$b_1(A) = 0.3, b_1(B) = 0.2, b_1(C) = 0.5$$

$$b_2(A) = 0.7, b_2(B) = 0.1, b_2(C) = 0.2$$

$$b_3(A) = 0.3, b_3(B) = 0.6, b_3(C) = 0.1$$

$$\pi = [0.4 \quad 0.5 \quad 0.1]$$



- Given a sequence of observations $O = \{ABC\}$, there are **27 possible** corresponding state sequences, and therefore the corresponding probability is

$$P(O|\lambda) = \sum_{i=1}^{27} P(O, S_i | \lambda) = \sum_{i=1}^{27} P(O|S_i, \lambda)P(S_i | \lambda), \quad S_i : \text{state sequence}$$

E.g. when $S_i = \{s_2 s_2 s_3\}$, $P(O|S_i, \lambda) = P(A|s_2)P(B|s_2)P(C|s_3) = 0.7 * 0.1 * 0.1 = 0.007$

$$P(S_i | \lambda) = P(s_2)P(s_2 | s_2)P(s_3 | s_2) = 0.5 * 0.7 * 0.2 = 0.07$$

Hidden Markov Model (cont.)

- Notation :
 - $O = \{o_1, o_2, o_3, \dots, o_T\}$: the observation (feature) sequence
 - $S = \{s_1, s_2, s_3, \dots, s_T\}$: the state sequence
 - λ : model, for HMM, $\lambda = \{A, B, \pi\}$
 - $P(O|\lambda)$: 用 model λ 計算 O 的機率值
 - $P(O|S, \lambda)$: 在 O 是 state sequence S 所產生的前提下, 用 model λ 計算 O 的機率值
 - $P(O, S|\lambda)$: 用 model λ 計算 $[O, S]$ 兩者同時成立的機率值
 - $P(S|O, \lambda)$: 在已知 O 的前提下, 用 model λ 計算 S 的機率值
- Useful formula
 - Bayesian Rule :

$$P(A|B) = \frac{P(A, B)}{P(B)} = \frac{P(B|A)P(A)}{P(B)} \rightarrow P(A|B, \lambda) = \frac{P(A, B|\lambda)}{P(B|\lambda)} = \frac{P(B|A, \lambda)P(A|\lambda)}{P(B|\lambda)}$$

λ : model describing the probability

$$P(A, B) = P(B|A)P(A) = P(A|B)P(B)$$

chain rule

Hidden Markov Model (cont.)

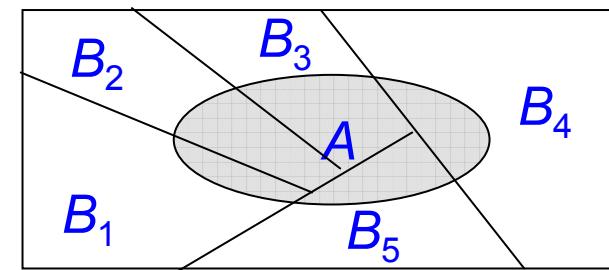
- Useful formula (Cont.):

marginal probability

$$P(A) = \begin{cases} \sum_{all B} P(A, B) = \sum_{all B} P(A|B)P(B), & \text{if } B \text{ is discrete and disjoint} \\ \int_B f(A, B)dB = \int_B f(A|B)f(B)dB, & \text{if } B \text{ is continuous} \end{cases}$$

if x_1, x_2, \dots, x_n are independent,

$$P(x_1, x_2, \dots, x_n) = P(x_1)P(x_2) \dots P(x_n)$$



$$E_z(q(z)) = \begin{cases} \sum_k P(z=k)q(k), & z: \text{discrete} \\ \int_z f_z(z)q(z)dz, & z: \text{continuous} \end{cases}$$

Expectation

Three Basic Problems for HMM

- Given an observation sequence $O=(o_1, o_2, \dots, o_T)$, and an HMM $\lambda=(S, A, B, \pi)$
 - **Problem 1:**
How to *efficiently* compute $P(O|\lambda)$?
 \Rightarrow ***Evaluation problem***
 - **Problem 2:**
How to choose an optimal state sequence $S=(s_1, s_2, \dots, s_T)$?
 \Rightarrow ***Decoding Problem***
 - **Problem 3:**
How to adjust the model parameter $\lambda=(A, B, \pi)$ to maximize $P(O|\lambda)$?
 \Rightarrow ***Learning / Training Problem***

Basic Problem 1 of HMM (cont.)

Given \mathbf{O} and λ , find $P(\mathbf{O}|\lambda) = \text{Prob}[\text{observing } \mathbf{O} \text{ given } \lambda]$

- Direct Evaluation

- Evaluating all possible state sequences of length T that generating observation sequence \mathbf{O}

$$P(\mathbf{o} | \lambda) = \sum_{\text{all } s} P(\mathbf{o}, s | \lambda) = \sum_{\text{all } s} P(\mathbf{o} | s, \lambda) P(s | \lambda)$$

- $P(s | \lambda)$: The probability of each path s
 - By Markov assumption (First-order HMM)

$$P(s | \lambda) = P(s_1 | \lambda) \prod_{t=2}^T P(s_t | s_1^{t-1}, \lambda)$$

By chain rule

$$\approx P(s_1 | \lambda) \prod_{t=2}^T P(s_t | s_{t-1}, \lambda)$$

By Markov assumption

$$= \pi_{s_1} a_{s_1 s_2} a_{s_2 s_3} \dots a_{s_{T-1} s_T}$$

Basic Problem 1 of HMM (cont.)

- Direct Evaluation (cont.)
 - $P(\mathbf{o} | \mathbf{s}, \lambda)$: The joint output probability along the path \mathbf{s}
 - By output-independent assumption
 - The probability that a particular observation symbol/vector is emitted at time t depends only on the state s_t and is conditionally independent of the past observations

$$\begin{aligned} P(\mathbf{o} | \mathbf{s}, \lambda) &= P(o_1 | s_1^T, \lambda) \\ &= P(o_1 | s_1^T, \lambda) \prod_{t=2}^T P(o_t | o_1^{t-1}, s_1^T, \lambda) \\ &\approx \prod_{t=1}^T P(o_t | s_t, \lambda) \quad \text{By output-independent assumption} \\ &= \prod_{t=1}^T b_{s_t}(\mathbf{o}_t) \end{aligned}$$

Basic Problem 1 of HMM (cont.)

- Direct Evaluation (Cont.)

$$P(\mathbf{o}_t | s_t, \lambda) = b_{s_t}(\mathbf{o}_t)$$

$$\begin{aligned} P(\mathbf{o} | \lambda) &= \sum_{\text{all } S} P(S | \lambda) P(\mathbf{o} | S, \lambda) \\ &= \sum_{\text{all } s} \left[\pi_{s_1} a_{s_1 s_2} a_{s_2 s_3} \dots a_{s_{T-1} s_T} \prod b_{s_1}(\mathbf{o}_1) b_{s_2}(\mathbf{o}_2) \dots b_{s_T}(\mathbf{o}_T) \right] \\ &= \sum_{s_1, s_2, \dots, s_T} \pi_{s_1} b_{s_1}(\mathbf{o}_1) a_{s_1 s_2} b_{s_2}(\mathbf{o}_2) \dots a_{s_{T-1} s_T} b_{s_T}(\mathbf{o}_T) \end{aligned}$$

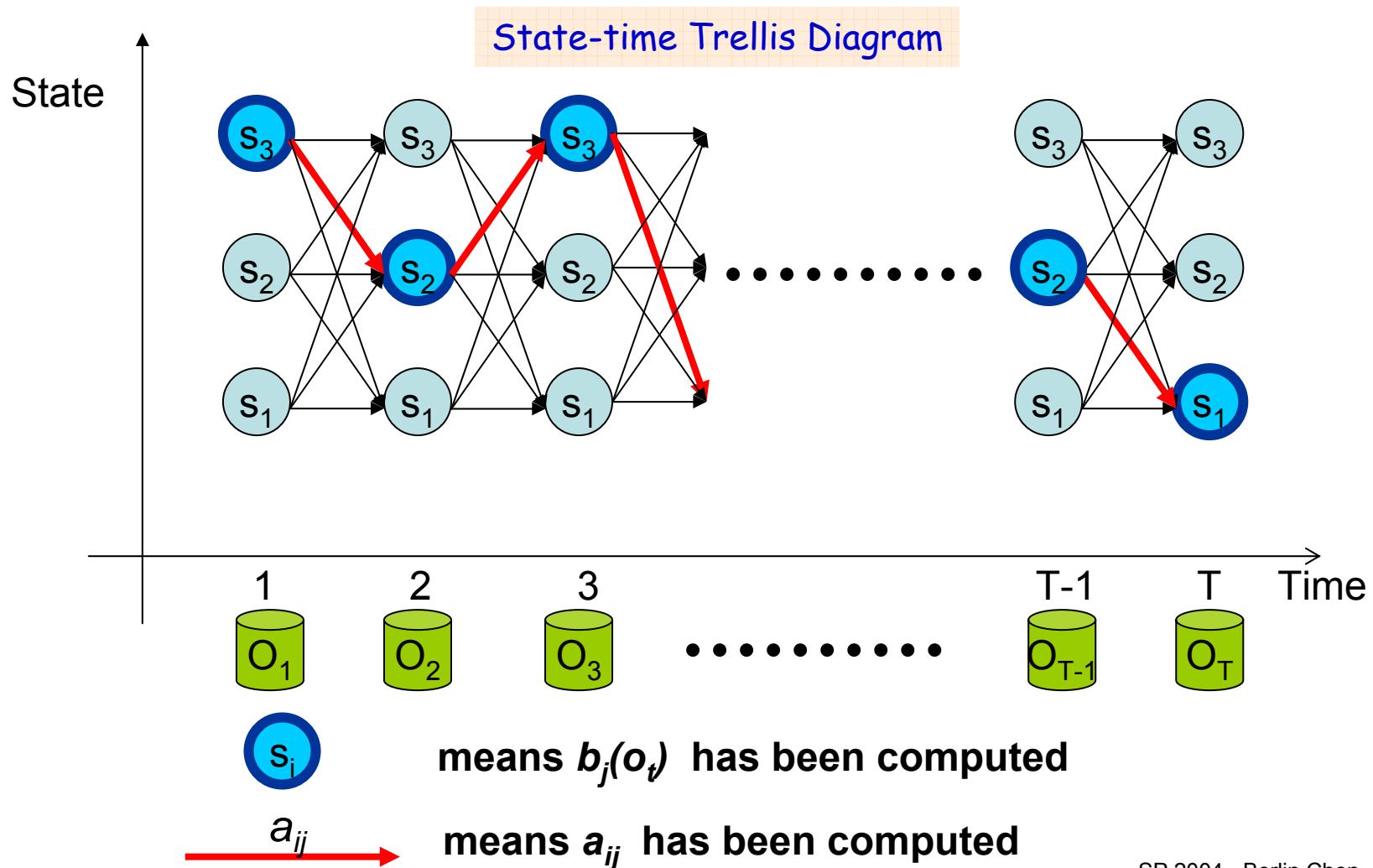
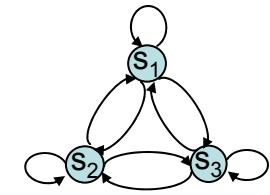
- Huge Computation Requirements: $O(N^T)$
 - Exponential computational complexity

Complexity : $(2T-1)N^T \text{ MUL} \approx 2TN^T, N^T - 1 \text{ ADD}$

- A more efficient algorithms can be used to evaluate $P(\mathbf{o} | \lambda)$
 - *Forward/Backward Procedure/Algorithm*

Basic Problem 1 of HMM (cont.)

- Direct Evaluation (Cont.)



Basic Problem 1 of HMM

- The Forward Procedure

- Base on the HMM assumptions, the calculation of $P(s_t | s_{t-1}, \lambda)$ and $P(o_t | s_t, \lambda)$ involves only s_{t-1} , s_t and o_t , so it is possible to compute the likelihood with recursion on t
- Forward variable : $\alpha_t(i) = P(o_1 o_2 \dots o_t, s_t = i | \lambda)$
 - The probability that the HMM is in state i at time t having generating partial observation $o_1 o_2 \dots o_t$

Basic Problem 1 of HMM

- The Forward Procedure (cont.)

- Algorithm

1. Initialization $\alpha_1(i) = \pi_i b_i(o_1)$, $1 \leq i \leq N$

2. Induction $\alpha_{t+1}(j) = \left[\sum_{i=1}^N \alpha_t(i) a_{ij} \right] b_j(o_{t+1})$, $1 \leq t \leq T-1, 1 \leq j \leq N$

3. Termination $P(O|\lambda) = \sum_{i=1}^N \alpha_T(i)$

– Complexity: $O(N^2T)$

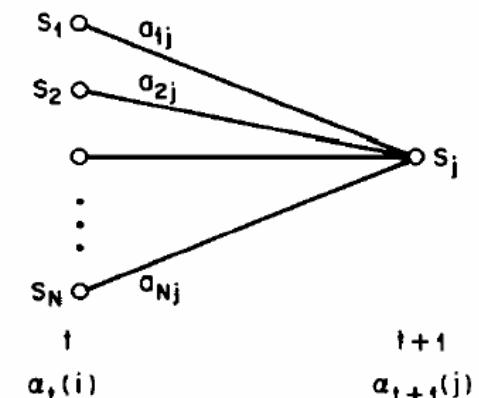
$$\text{MUL : } N(N+1)(T-1) + N \approx N^2T$$

$$\text{ADD : } (N-1)N(T-1) + (N-1) \approx N^2T$$

- Based on the lattice (trellis) structure

– Computed in a *time-synchronous* fashion from *left-to-right*, where each cell for time t is completely computed before proceeding to time $t+1$

- All state sequences, regardless how long previously, merge to N nodes (states) at each time instance t



Basic Problem 1 of HMM

- The Forward Procedure (cont.)

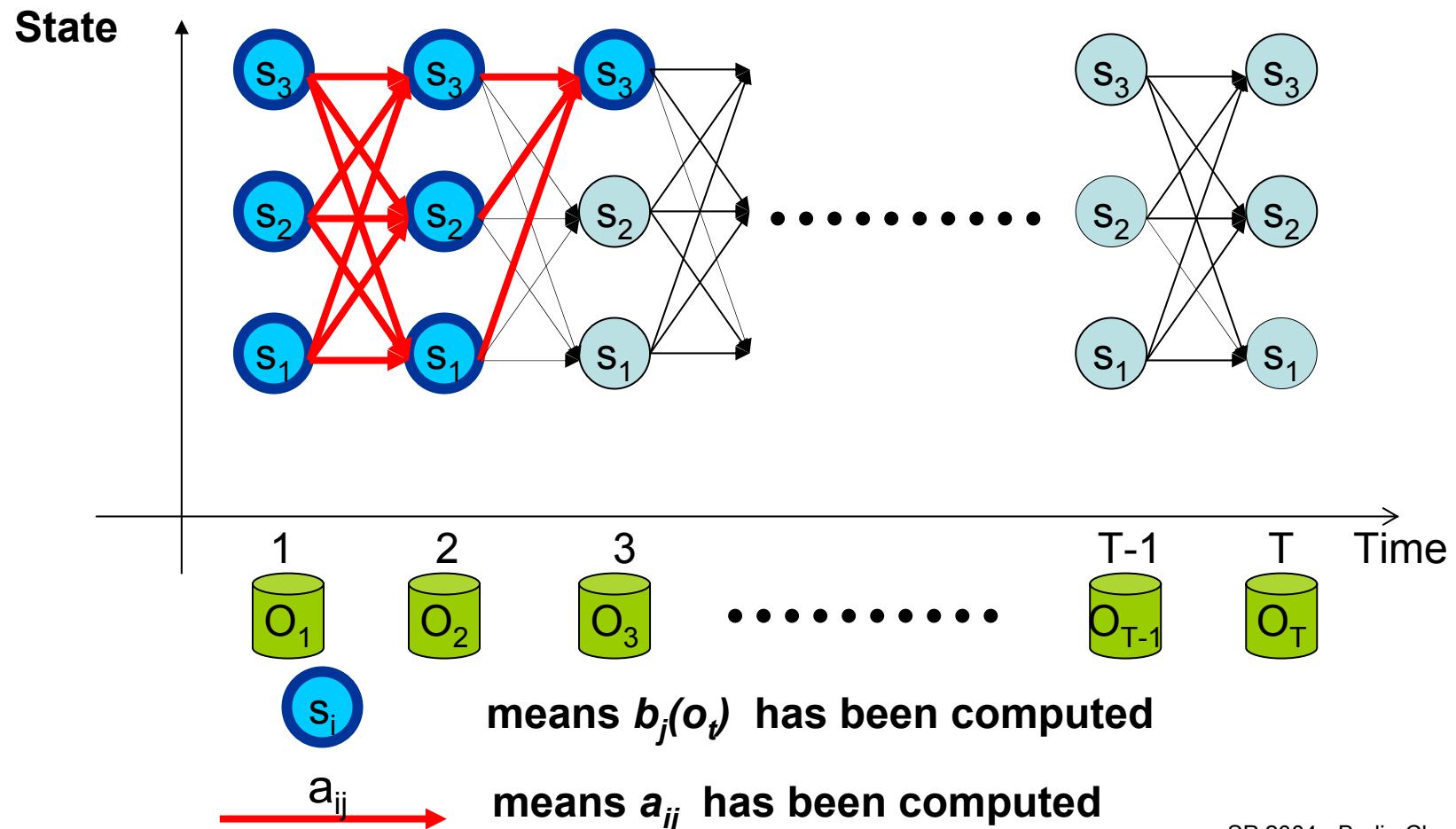
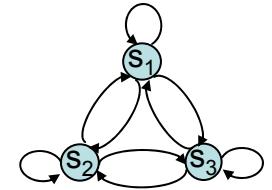
$$\begin{aligned}
 \alpha_t(j) &= P(o_1 o_2 \dots o_t, s_t = j | \lambda) \\
 &= P(o_1 o_2 \dots o_t | s_t = j, \lambda) P(s_t = j | \lambda) \\
 &= P(o_1 o_2 \dots o_{t-1} | s_t = j, \lambda) P(o_t | s_t = j, \lambda) P(s_t = j | \lambda) \\
 &= P(o_1 o_2 \dots o_{t-1}, s_t = j | \lambda) P(o_t | s_t = j, \lambda) \\
 &= P(o_1 o_2 \dots o_{t-1}, s_t = j | \lambda) b_j(o_t) \\
 &= \left[\sum_{i=1}^N P(o_1 o_2 \dots o_{t-1}, s_{t-1} = i, s_t = j | \lambda) \right] b_j(o_t) \\
 &= \left[\sum_{i=1}^N P(o_1 o_2 \dots o_{t-1}, s_{t-1} = i | \lambda) P(s_t = j | o_1 o_2 \dots o_{t-1}, s_{t-1} = i, \lambda) \right] b_j(o_t) \\
 &= \left[\sum_{i=1}^N P(o_1 o_2 \dots o_{t-1}, s_{t-1} = i | \lambda) P(s_t = j | s_{t-1} = i, \lambda) \right] b_j(o_t) \\
 &= \left[\sum_{i=1}^N \alpha_{t-1}(i) a_{ij} \right] b_j(o_t)
 \end{aligned}$$

$P(A, B) = P(B|A)P(A)$
 output independent assumption
 $P(B|A)P(A) = P(A, B)$
 $P(o_t | s_t = j, \lambda) = b_j(o_t)$
 $P(A) \sum_{all B} P(A, B)$
 first-order Markov assumption

Basic Problem 1 of HMM

- The Forward Procedure (cont.)

- $\alpha_3(3) = P(o_1, o_2, o_3, s_3=3 | \lambda)$
 $= [\alpha_2(1)^* a_{13} + \alpha_2(2)^* a_{23} + \alpha_2(3)^* a_{33}] b_3(o_3)$



Basic Problem 1 of HMM

- The Forward Procedure (cont.)

- A three-state Hidden Markov Model for the *Dow Jones Industrial average*

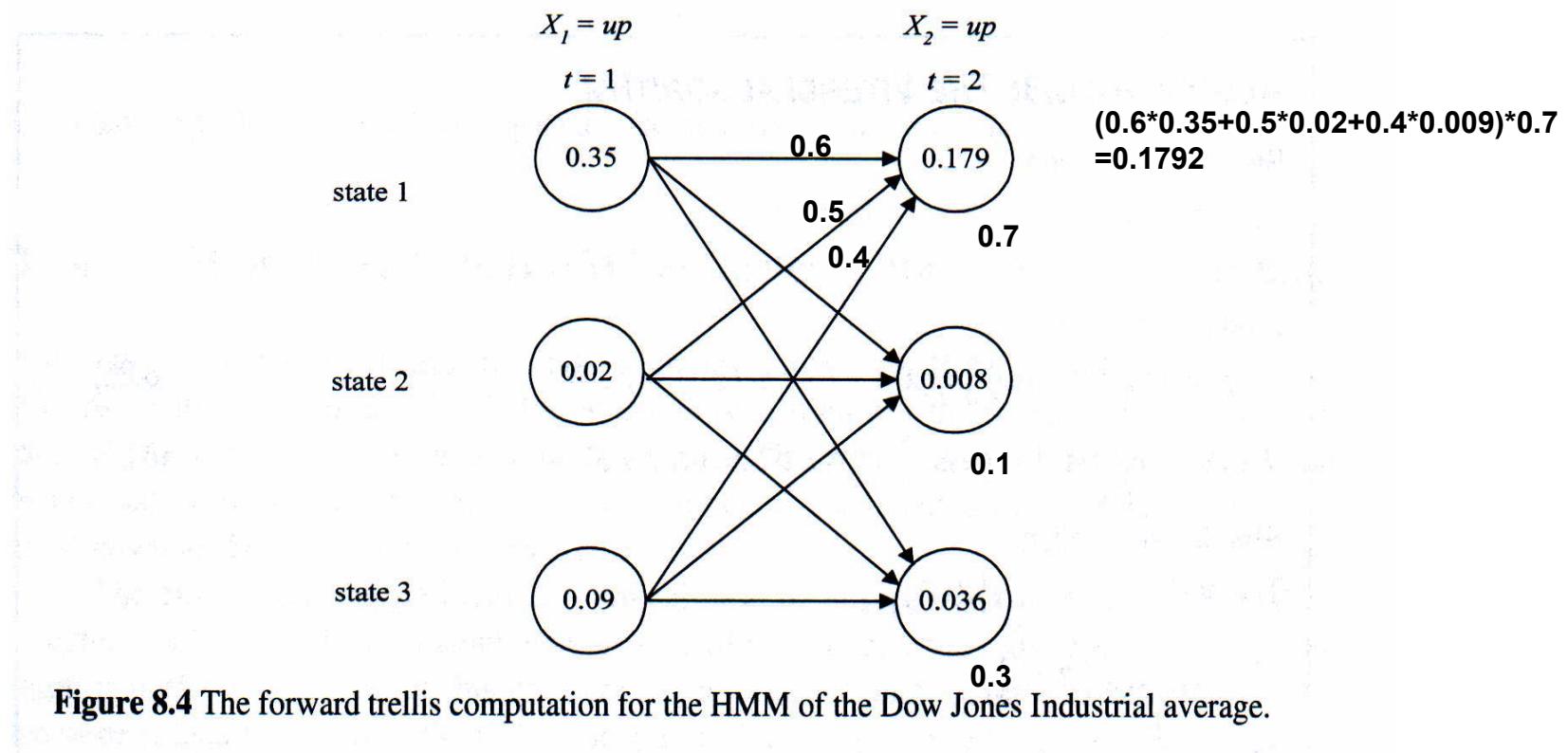


Figure 8.4 The forward trellis computation for the HMM of the Dow Jones Industrial average.

Basic Problem 1 of HMM

- The Backward Procedure

- Backward variable : $\beta_t(i) = P(\mathbf{o}_{t+1}, \mathbf{o}_{t+2}, \dots, \mathbf{o}_T | s_t = i, \lambda)$

1. Initialization : $\beta_T(i) = 1, 1 \leq i \leq N$

2. Induction : $\beta_t(i) = \sum_{j=1}^N a_{ij} b_j(\mathbf{o}_{t+1}) \beta_{t+1}(j), 1 \leq t \leq T-1, 1 \leq i \leq N$

3. Termination : $P(\mathbf{O}|\lambda) = \sum_{j=1}^N \pi_j b_j(\mathbf{o}_1) \beta_1(j)$

Complexity MUL: $2N^2(T-1) + 2N \approx N^2T$;

ADD: $(N-1)N(T-1) + N \approx N^2T$

Basic Problem 1 of HMM

- Backward Procedure (cont.)

- Why $P(\mathbf{O}, s_t = i | \lambda) = \alpha_t(i) \beta_t(i)$?

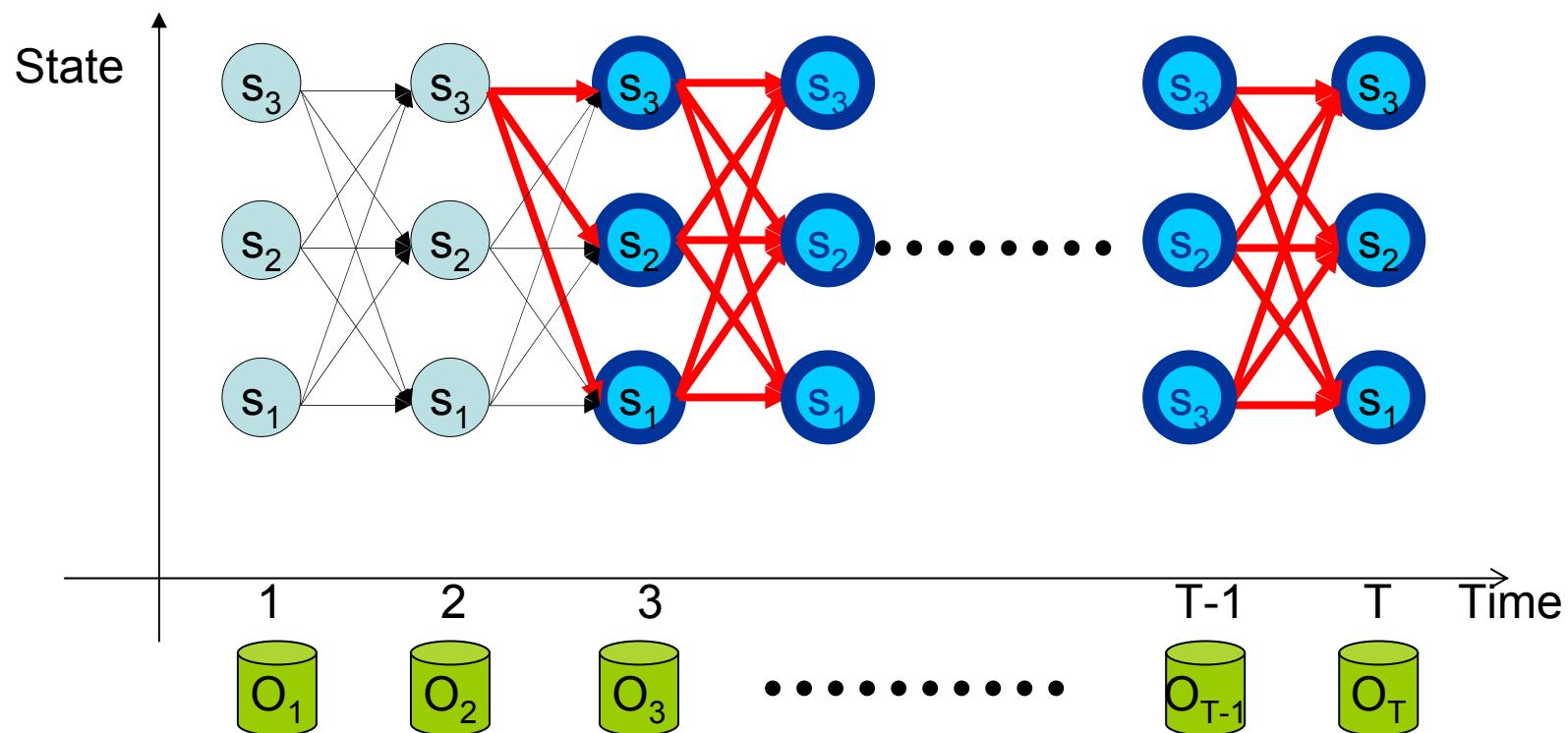
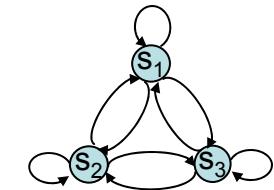
$$\begin{aligned} & \alpha_t(i) \beta_t(i) \\ &= P(\mathbf{o}_1, \mathbf{o}_2, \dots, \mathbf{o}_t, s_t = i | \lambda) \cdot P(\mathbf{o}_{t+1}, \mathbf{o}_{t+2}, \dots, \mathbf{o}_T | s_t = i, \lambda) \\ &= P(\mathbf{o}_1, \mathbf{o}_2, \dots, \mathbf{o}_t | s_t = i, \lambda) P(s_t = i | \lambda) P(\mathbf{o}_{t+1}, \mathbf{o}_{t+2}, \dots, \mathbf{o}_T | s_t = i, \lambda) \\ &= P(\mathbf{o}_1, \dots, \mathbf{o}_t, \dots, \mathbf{o}_T | s_t = i, \lambda) P(s_t = i | \lambda) \\ &= P(\mathbf{o}_1, \dots, \mathbf{o}_t, \dots, \mathbf{o}_T, s_t = i | \lambda) \\ &= P(\mathbf{O}, s_t = i | \lambda) \end{aligned}$$

- $P(\mathbf{O} | \lambda) = \sum_{i=1}^N P(\mathbf{O}, s_t = i | \lambda) = \sum_{i=1}^N \alpha_t(i) \beta_t(i)$

Basic Problem 1 of HMM

- The Backward Procedure (cont.)

- $$\begin{aligned}\beta_2(3) &= P(o_3, o_4, \dots, o_T | s_2 = 3, \lambda) \\ &= a_{31} * b_1(o_3) * \beta_3(1) + a_{32} * b_2(o_3) * \beta_3(2) + a_{33} * b_3(o_3) * \beta_3(3)\end{aligned}$$



Basic Problem 2 of HMM

How to choose an optimal state sequence $S=(s_1, s_2, \dots, s_T)$?

- The first optimal criterion: **Choose the states s_t are individually most likely at each time t**

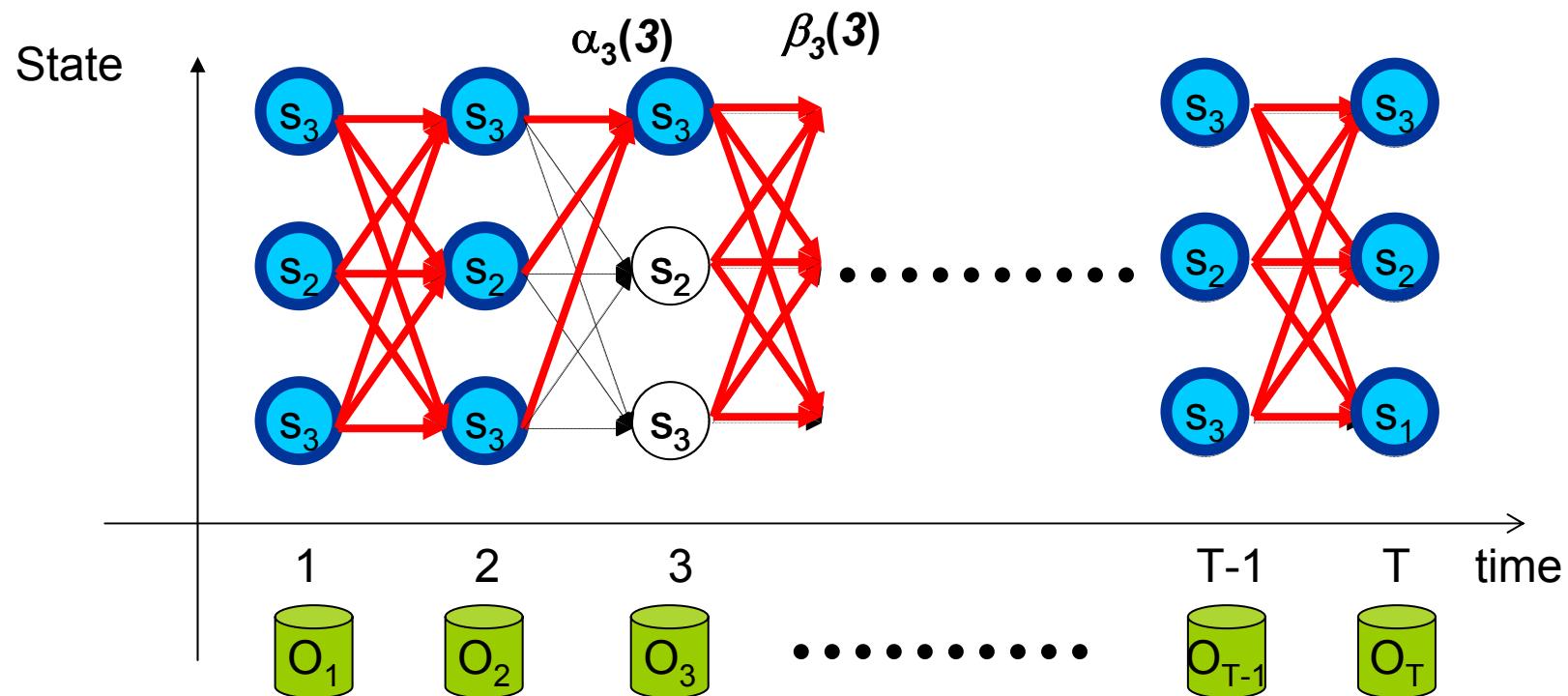
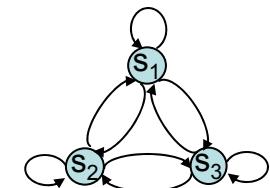
Define a posteriori probability variable $\gamma_t(i) = P(s_t = i | O, \lambda)$

$$\gamma_t(i) = \frac{P(s_t = i, O | \lambda)}{P(O | \lambda)} = \frac{P(s_t = i, O | \lambda)}{\sum_{m=1}^N P(s_t = m, O | \lambda)} = \frac{\alpha_t(i) \beta_t(i)}{\sum_{m=1}^N \alpha_t(m) \beta_t(m)}$$

- Solution : $s_t^* = \arg_i \max [\gamma_t(i)], 1 \leq t \leq T$
 - Problem: maximizing the probability at each time t individually
 $S^* = s_1^* s_2^* \dots s_T^*$ may not be a valid sequence (e.g. $a_{s_t^* s_{t+1}^*} = 0$)

Basic Problem 2 of HMM (cont.)

- $P(s_3 = 3, O | \lambda) = \alpha_3(3) * \beta_3(3)$



Basic Problem 2 of HMM

- The Viterbi Algorithm

- The second optimal criterion: The Viterbi algorithm can be regarded as the dynamic programming algorithm applied to the HMM or as a modified forward algorithm
 - Instead of summing up probabilities from different paths coming to the same destination state, the Viterbi algorithm picks and remembers the best path
 - Find a single optimal state sequence $S=(s_1, s_2, \dots, s_T)$
 - How to find the second, third, etc., optimal state sequences (difficult?)
 - The Viterbi algorithm also can be illustrated in a trellis framework similar to the one for the forward algorithm
 - State-time trellis diagram

Basic Problem 2 of HMM

- The Viterbi Algorithm (cont.)

- **Algorithm**

Find a best state sequence $S = (s_1, s_2, \dots, s_T)$ for a given observation $O = (o_1, o_2, \dots, o_T)$?

Define a new variable

$$\delta_t(i) = \max_{s_1, s_2, \dots, s_{t-1}} P[s_1, s_2, \dots, s_{t-1}, s_t = i, o_1, o_2, \dots, o_t | \lambda]$$

= the best score along a single path at time t , which accounts for the first t observations and ends in state i

By induction $\therefore \delta_{t+1}(j) = \left[\max_{1 \leq i \leq N} \delta_t(i) a_{ij} \right] b_j(o_{t+1})$

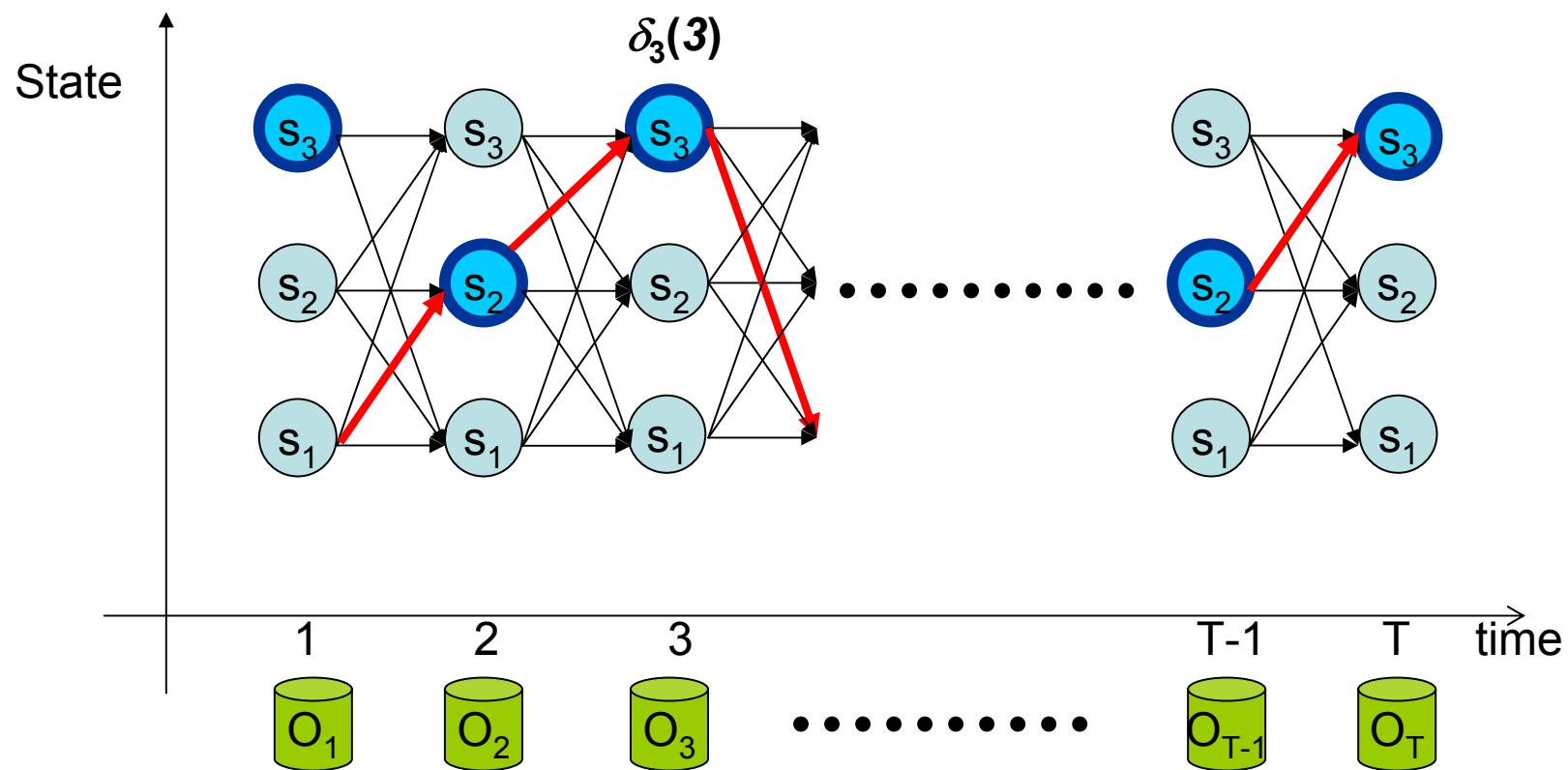
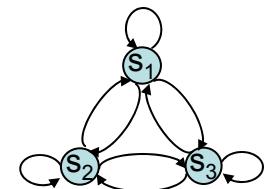
$$\psi_{t+1}(j) = \arg \max_{1 \leq i \leq N} \delta_t(i) a_{ij} \text{ For backtracing}$$

We can backtrace from $s_T^* = \arg \max_{1 \leq i \leq N} \delta_T(i)$

- Complexity: $O(N^2 T)$

Basic Problem 2 of HMM

- The Viterbi Algorithm (cont.)



Basic Problem 2 of HMM

- The Viterbi Algorithm (cont.)

- A three-state Hidden Markov Model for the *Dow Jones Industrial average*

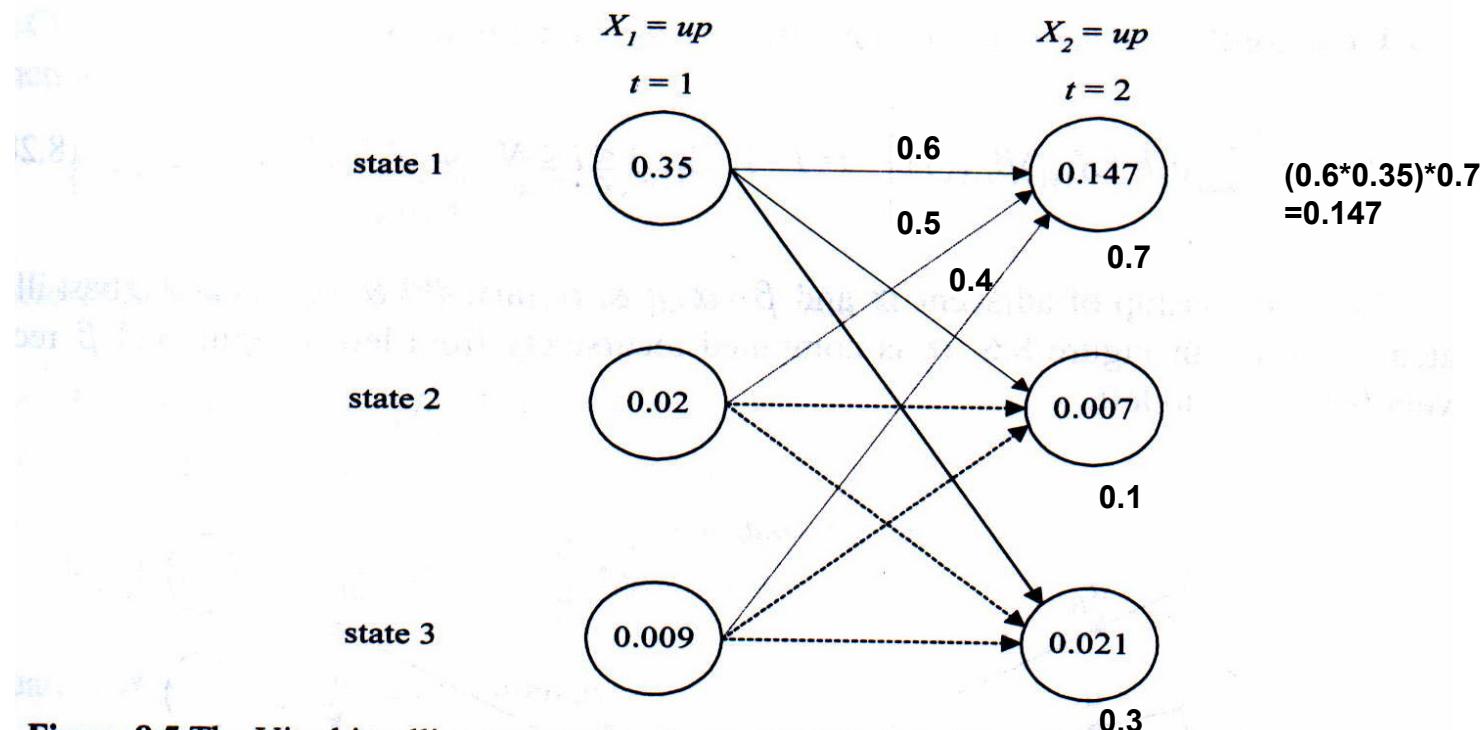


Figure 8.5 The Viterbi trellis computation for the HMM of the Dow Jones Industrial average.

Basic Problem 2 of HMM

- The Viterbi Algorithm (cont.)

- Algorithm in the **logarithmic form**

Find a best state sequence $S = (s_1, s_2, \dots, s_T)$ for a given observation $O = (o_1, o_2, \dots, o_T)$?

Define a new variable

$$\delta_t(i) = \max_{s_1, s_2, \dots, s_{t-1}} \log P[s_1, s_2, \dots, s_{t-1}, s_t = i, o_1, o_2, \dots, o_t | \lambda]$$

= the best score along a single path at time t , which accounts for the first t observations and ends in state i

By induction: $\delta_{t+1}(j) = \left[\max_{1 \leq i \leq N} (\delta_t(i) + \log a_{ij}) \right] + \log b_j(o_{t+1})$

$$\psi_{t+1}(j) = \arg \max_{1 \leq i \leq N} (\delta_t(i) + \log a_{ij}) \dots \text{For backtracing}$$

We can backtrace from $s_T^* = \arg \max_{1 \leq i \leq N} \delta_T(i)$

Homework-1

- A three-state Hidden Markov Model for the *Dow Jones Industrial average*

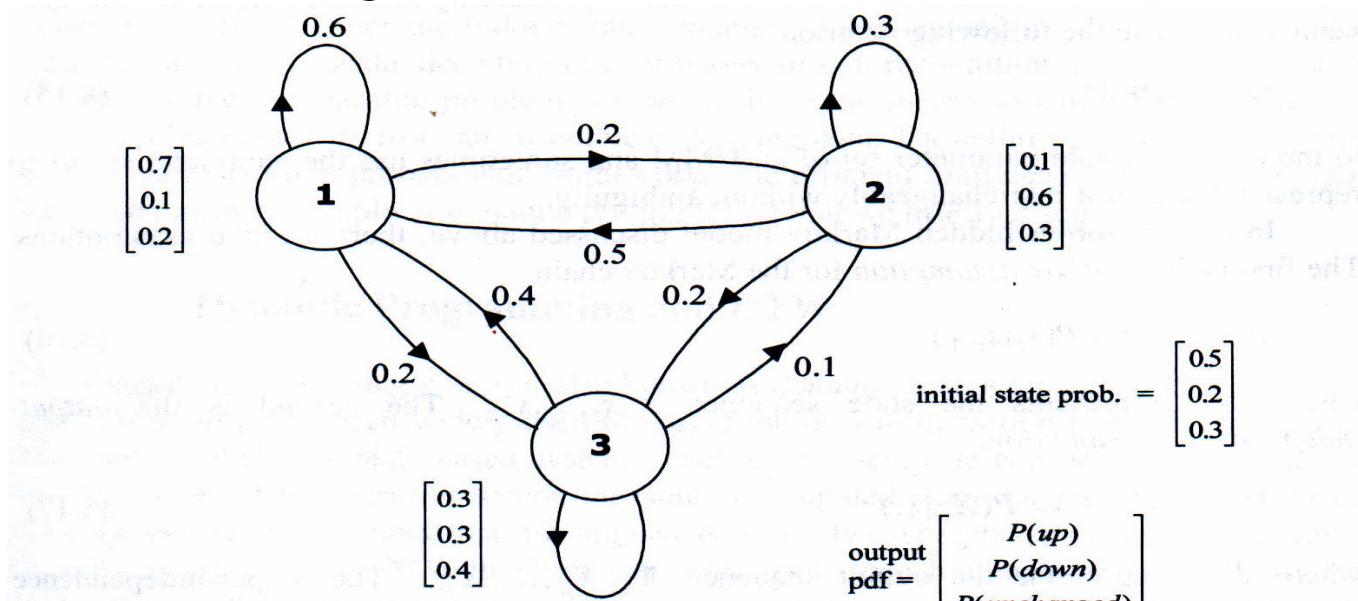
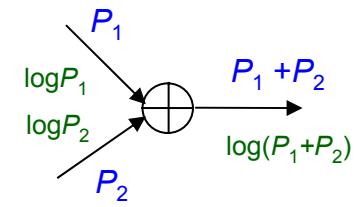


Figure 8.2 A hidden Markov model for the Dow Jones Industrial average. The three states no longer have deterministic meanings as in the Markov chain illustrated in Figure 8.1.

- Find the probability:
 $P(\text{up}, \text{up}, \text{unchanged}, \text{down}, \text{unchanged}, \text{down}, \text{up} | \lambda)$
- Find the optimal state sequence of the model which generates the observation sequence: (up, up, unchanged, down, unchanged, down, up)

Probability Addition in F-B Algorithm

- In Forward-backward algorithm, operations usually implemented in logarithmic domain
- Assume that we want to add P_1 and P_2



if $P_1 \geq P_2$

$$\log_b(P_1 + P_2) = \log_b P_1 + \log_b \left(1 + b^{\log_b P_2 - \log_b P_1} \right)$$

else

$$\log_b(P_1 + P_2) = \log_b P_2 + \log_b \left(1 + b^{\log_b P_1 - \log_b P_2} \right)$$

The values of $\log_b(1 + b^x)$ can be saved in a table to speedup the operations

Probability Addition in F-B Algorithm (cont.)

- An example code

```
#define LZERO (-1.0E10) // ~log(0)
#define LSMALL (-0.5E10) // log values < LSMALL are set to LZERO
#define minLogExp -log(-LZERO)

double LogAdd(double x, double y)
{
    double temp,diff,z;
    if (x<y)
    {
        temp = x; x = y; y = temp;
    }
    diff = y-x; //notice that diff <= 0
    if (diff<minLogExp) // if y' is far smaller than x'
        return (x<LSMALL) ? LZERO:x;
    else
    {
        z = exp(diff);
        return x+log(1.0+z);
    }
}
```

Basic Problem 3 of HMM

Intuitive View

- How to adjust (re-estimate) the model parameter $\lambda = (\mathbf{A}, \mathbf{B}, \pi)$ to maximize $P(\mathbf{O}|\lambda)$?
 - The most difficult of the three problems, because there is no known analytical method that maximizes the joint probability of the training data in a close form
 - The **data is incomplete** because of the hidden state sequences
 - Well-solved by the *Baum-Welch* (known as *forward-backward*) algorithm and *EM (Expectation-Maximization)* algorithm
 - Iterative update and improvement

Basic Problem 3 of HMM

Intuitive View (cont.)

- Relation between the forward and backward variables

$$\alpha_t(i) = P(\mathbf{o}_1, \mathbf{o}_2, \dots, \mathbf{o}_t, s_t = i | \lambda)$$

$$= \left[\sum_{j=1}^N \alpha_{t-1}(j) a_{ji} \right] b_i(\mathbf{o}_t)$$

$$\beta_t(i) = P(\mathbf{o}_{t+1}, \mathbf{o}_{t+2}, \dots, \mathbf{o}_T | s_t = i, \lambda)$$

$$= \sum_{j=1}^N \beta_{t+1}(j) b_j(\mathbf{o}_{t+1}) a_{ij}$$

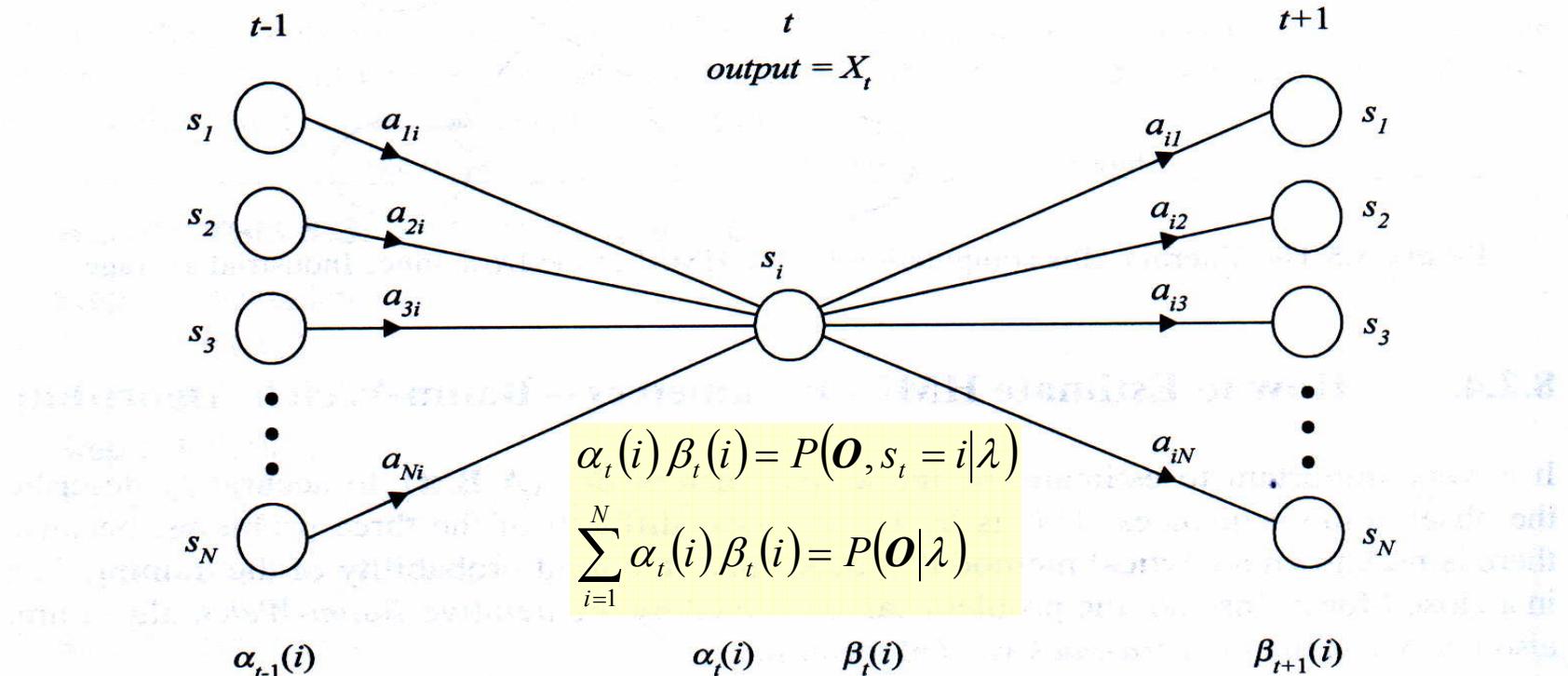


Figure 8.6 The relationship of α_{t-1} and α_t and β_t and β_{t+1} in the forward-backward algorithm.

Basic Problem 3 of HMM

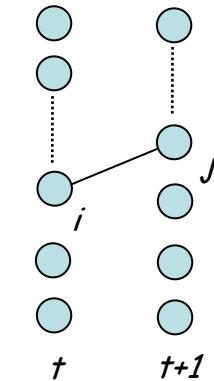
Intuitive View (cont.)

- Define a new variable:

$$\xi_t(i, j) = P(s_t = i, s_{t+1} = j | \mathbf{O}, \lambda)$$

- Probability being at state i at time t and at state j at time $t+1$

$$\begin{aligned}\xi_t(i, j) &= \frac{P(s_t = i, s_{t+1} = j, \mathbf{O} | \lambda)}{P(\mathbf{O} | \lambda)} \\ &= \frac{\alpha_t(i) a_{ij} b_j(\mathbf{o}_{t+1}) \beta_{t+1}(j)}{P(\mathbf{O} | \lambda)} = \frac{\alpha_t(i) a_{ij} b_j(\mathbf{o}_{t+1}) \beta_{t+1}(j)}{\sum_{m=1}^N \sum_{n=1}^N \alpha_t(m) a_{mn} b_n(\mathbf{o}_{t+1}) \beta_{t+1}(n)}\end{aligned}$$



- Recall the posterior probability variable:

$$\gamma_t(i) = P(s_t = i | \mathbf{O}, \lambda)$$

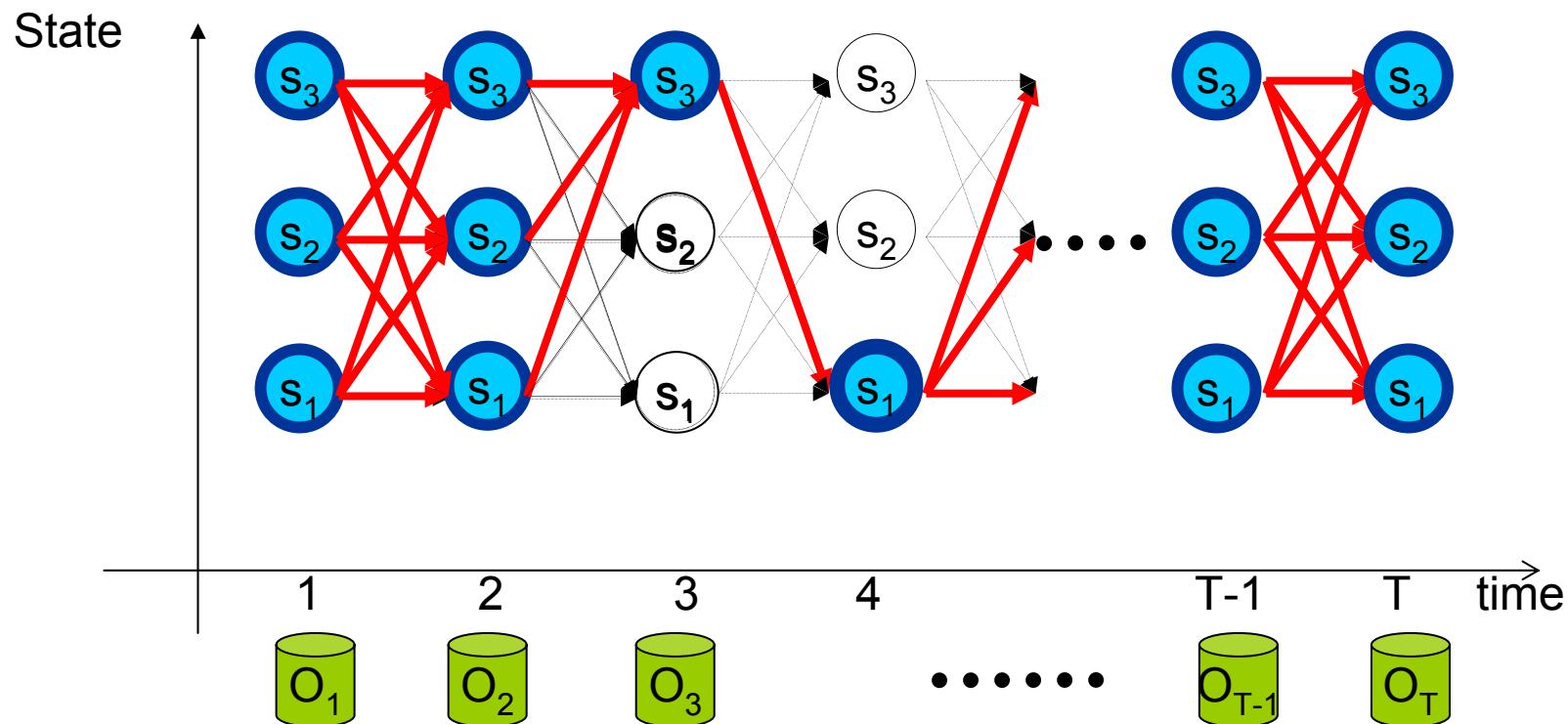
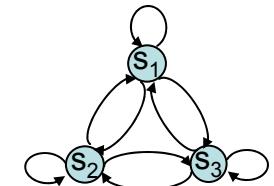
Note : $\gamma_t(i)$ also can be represented as $\frac{\alpha_t(i) \beta_t(i)}{\sum_{m=1}^N \alpha_t(m) \beta_t(m)}$

$$\gamma_t(i) = \sum_{j=1}^N P(s_t = i, s_{t+1} = j | \mathbf{O}, \lambda) = \sum_{j=1}^N \xi_t(i, j) \quad (\text{for } t < T)$$

Basic Problem 3 of HMM

Intuitive View (cont.)

- $P(s_3 = 3, s_4 = 1, O | \lambda) = \alpha_3(3) * a_{31} * b_1(o_4) * \beta_1(4)$



Basic Problem 3 of HMM

Intuitive View (cont.)

- $\xi_t(i, j) = P(s_t = i, s_{t+1} = j | \mathbf{O}, \lambda)$

$\sum_{t=1}^{T-1} \xi_t(i, j) = \text{expected number of transitions from state } i \text{ to state } j \text{ in } \mathbf{O}$

- $\gamma_t(i) = P(s_t = i | \mathbf{O}, \lambda)$

$\sum_{t=1}^{T-1} \gamma_t(i) = \sum_{t=1}^{T-1} \sum_{j=1}^N \xi_t(i, j) = \text{expected number of transitions from state } i \text{ in } \mathbf{O}$

- A set of reasonable re-estimation formula for $\{\mathbf{A}, \pi\}$ is

$$\begin{aligned}\bar{\pi}_i &= \text{expected frequency (number of times) in state } i \text{ at time } t=1 \\ &= \gamma_1(i)\end{aligned}$$

$$\bar{a}_{ij} = \frac{\text{expected number of transition from state } i \text{ to state } j}{\text{expected number of transition from state } i} = \frac{\sum_{t=1}^{T-1} \xi_t(i, j)}{\sum_{t=1}^{T-1} \gamma_t(i)}$$

Formulae for Single Training Utterance

Basic Problem 3 of HMM

Intuitive View (cont.)

- A set of reasonable re-estimation formula for $\{B\}$ is
 - For discrete and finite observation $b_j(v_k) = P(o_t = v_k | s_t = j)$

$$\bar{b}_j(v_k) = \bar{P}(o = v_k | s = j) = \frac{\text{expected number of times in state } j \text{ and observing symbol } v_k}{\text{expected number of times in state } j} = \frac{\sum_{t=1}^T \gamma_t(j) \text{ such that } o=v_k}{\sum_{t=1}^T \gamma_t(j)}$$

- For continuous and infinite observation $b_j(v) = f_{o|s}(o_t = v | s_t = j)$,

$$\bar{b}_j(v) = \sum_{k=1}^M \bar{c}_{jk} N(v; \bar{\mu}_{jk}, \bar{\Sigma}_{jk}) = \sum_{k=1}^M \bar{c}_{jk} \left(\frac{1}{(\sqrt{2\pi})^L |\bar{\Sigma}_{jk}|^{1/2}} \exp \left(-\frac{1}{2} (v - \bar{\mu}_{jk})^t \bar{\Sigma}_{jk}^{-1} (v - \bar{\mu}_{jk}) \right) \right)$$

Modeled as a mixture of multivariate Gaussian distributions

Basic Problem 3 of HMM

Intuitive View (cont.)

- For continuous and infinite observation (Cont.)

- Define a new variable $\gamma_t(j, k)$

– $\gamma_t(j, k)$ is the probability of being in state j at time t with the k -th mixture component accounting for \mathbf{o}_t

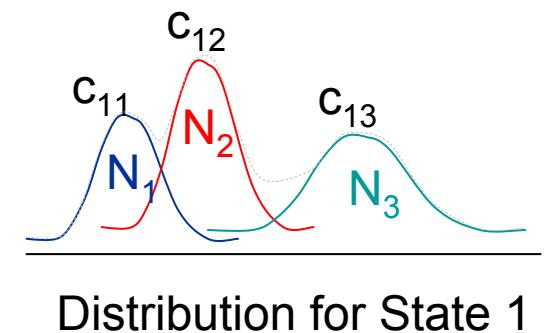
$$\begin{aligned}\gamma_t(j, k) &= P(s_t = j, m_t = k | \mathbf{O}, \lambda) \\ &= P(s_t = j | \mathbf{O}, \lambda)P(m_t = k | s_t = j, \mathbf{O}, \lambda) \\ &= \gamma_t(j)P(m_t = k | s_t = j, \mathbf{O}, \lambda) \\ &= \gamma_t(j) \frac{P(m_t = k, \mathbf{O} | s_t = j, \lambda)}{P(\mathbf{O} | s_t = j, \lambda)} \\ &= \gamma_t(j) \frac{P(m_t = k | s_t = j, \lambda)P(\mathbf{O} | s_t = j, m_t = k, \lambda)}{P(\mathbf{O} | s_t = j, \lambda)}\end{aligned}$$

..... (observation - independent assumption is applied)

$$= \gamma_t(j) \frac{P(m_t = k | s_t = j, \lambda)P(\mathbf{o}_t | s_t = j, m_t = k, \lambda)}{P(\mathbf{o}_t | s_t = j, \lambda)}$$

$$= \left[\frac{\alpha_t(j)\beta_t(j)}{\sum_{s=1}^N \alpha_t(s)\beta_t(s)} \right] \left[\frac{c_{jk} N(\mathbf{o}_t; \boldsymbol{\mu}_{jk}, \boldsymbol{\Sigma}_{jk})}{\sum_{m=1}^M c_{jm} N(\mathbf{o}_t; \boldsymbol{\mu}_{jm}, \boldsymbol{\Sigma}_{jm})} \right]$$

$$p(A|B) = \frac{p(A, B)}{P(B)}$$



$$\text{Note: } \gamma_t(j) = \sum_{m=1}^M \gamma_t(j, m)$$

Basic Problem 3 of HMM

Intuitive View (cont.)

- For continuous and infinite observation (Cont.)

$$\bar{c}_{jk} = \frac{\text{expected number of times in state } j \text{ and mixture } k}{\text{expected number of times in state } j} = \frac{\sum_{t=1}^T \gamma_t(j, k)}{\sum_{t=1}^T \sum_{m=1}^M \gamma_t(j, m)}$$

$$\bar{\mu}_{jk} = \text{weighted average (mean) of observations at state } j \text{ and mixture } k = \frac{\sum_{t=1}^T \gamma_t(j, k) \cdot o_t}{\sum_{t=1}^T \gamma_t(j, k)}$$

$$\begin{aligned}\bar{\Sigma}_{jk} &= \text{weighted covariance of observations at state } j \text{ and mixture } k \\ &= \frac{\sum_{t=1}^T \gamma_t(j, k) \cdot (o_t - \bar{\mu}_{jk})(o_t - \bar{\mu}_{jk})^t}{\sum_{t=1}^T \gamma_t(j, k)}\end{aligned}$$

Basic Problem 3 of HMM

Intuitive View (cont.)

- For continuous and infinite observation (Cont.)

$$\bar{\pi}_i = \text{expected frequency (number of times) in state } i \text{ at time } (t=1) = \frac{1}{L} \sum_{l=1}^L \gamma_1^l(i)$$

$$\bar{a}_{ij} = \frac{\text{expected number of transition from state } i \text{ to state } j}{\text{expected number of transition from state } i} = \frac{\sum_{l=1}^L \sum_{t=1}^{T-1} \xi_t^l(i,j)}{\sum_{l=1}^L \sum_{t=1}^{T-1} \gamma_t^l(i)}$$

$$\bar{c}_{jk} = \frac{\text{expected number of times in state } j \text{ and mixture } k}{\text{expected number of times in state } j} = \frac{\sum_{l=1}^L \sum_{t=1}^T \gamma_t^l(j,k)}{\sum_{l=1}^L \sum_{t=1}^T \sum_{m=1}^M \gamma_t^l(j,m)}$$

$$\bar{\mu}_{jk} = \text{weighted average (mean) of observations at state } j \text{ and mixture } k = \frac{\sum_{l=1}^L \sum_{t=1}^T \gamma_t^l(j,k) \cdot \mathbf{o}_t}{\sum_{l=1}^L \sum_{t=1}^T \gamma_t^l(j,k)}$$

$\bar{\Sigma}_{jk}$ = weighted covariance of observations at state j and mixture k

$$= \frac{\sum_{l=1}^L \sum_{t=1}^T \gamma_t^l(j,k) \cdot (\mathbf{o}_t - \bar{\mu}_{jk}) (\mathbf{o}_t - \bar{\mu}_{jk})^T}{\sum_{l=1}^L \sum_{t=1}^T \gamma_t^l(j,k)}$$

Formulae for Multiple (L) Training Utterances

Basic Problem 3 of HMM

Intuitive View (cont.)

- For discrete and finite observation (cont.)

$$\bar{\pi}_i = \text{expected frequency (number of times) in state } i \text{ at time } (t=1) = \frac{1}{L} \sum_{l=1}^L \gamma_1^l(i)$$

$$\bar{a}_{ij} = \frac{\text{expected number of transition from state } i \text{ to state } j}{\text{expected number of transition from state } i} = \frac{\sum_{l=1}^L \sum_{t=1}^{T-1} \xi_t^l(i,j)}{\sum_{l=1}^L \sum_{t=1}^{T-1} \gamma_t^l(i)}$$

$$\bar{b}_j(\mathbf{v}_k) = \bar{P}(\mathbf{o} = \mathbf{v}_k | s = j) = \frac{\text{expected number of times in state } j \text{ and observing symbol } \mathbf{v}_k}{\text{expected number of times in state } j} = \frac{\sum_{l=1}^L \sum_{t=1}^T \gamma_t^l(j) \text{ such that } \mathbf{o} = \mathbf{v}_k}{\sum_{l=1}^L \sum_{t=1}^T \gamma_t^l(j)}$$

Semicontinuous HMMs

- The HMM state mixture density functions are tied together across all the models to form a set of shared kernels

- The semicontinuous or tied-mixture HMM

$$b_j(\mathbf{o}) = \sum_{k=1}^M b_j(k) f(\mathbf{o}|v_k) = \sum_{k=1}^M b_j(k) N(\mathbf{o}, \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

state output

Probability of state j

k-th mixture weight

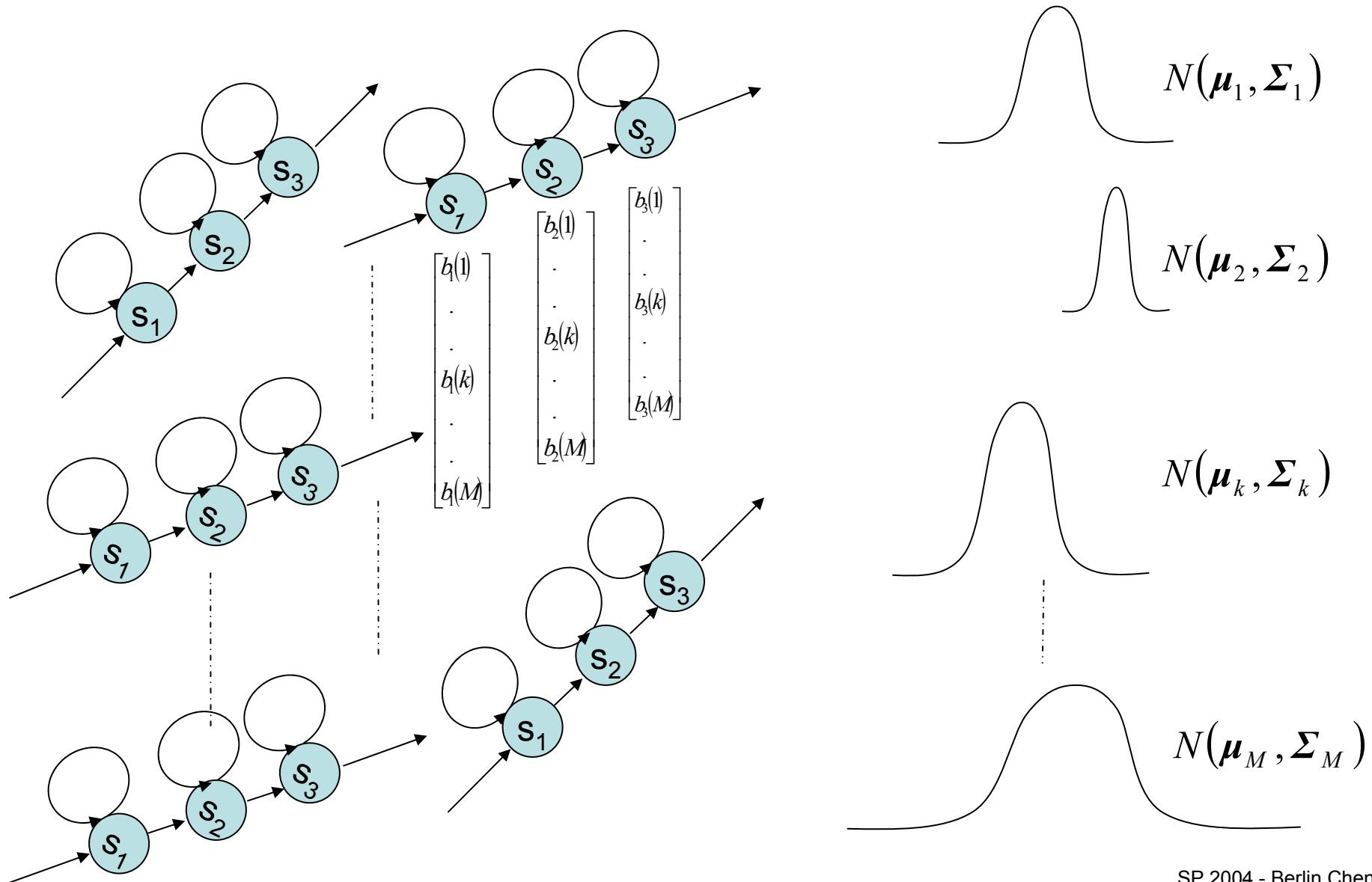
t of state /

(discrete, model-dependent)

k -th mixture density function or k -th codeword
(shared across HMMs, M is very large)

- A combination of the discrete HMM and the continuous HMM
 - A combination of *discrete* model-dependent weight coefficients and *continuous* model-independent codebook probability density functions
 - Because M is large, we can simply use the L most significant values $f(o|v_k)$
 - Experience showed that L is 1~3% of M is adequate
 - Partial tying of $f(o|v_k)$ for different phonetic class

Semicontinuous HMMs (cont.)



Initialization of HMM

- A good initialization of HMM training :

Segmental K-Means Segmentation into States

- Assume that we have a training set of observations and an initial estimate of all model parameters
- Step 1 : The set of training observation sequences is segmented into states, based on the initial model (finding the optimal state sequence by *Viterbi* Algorithm)
- Step 2 :
 - For discrete density HMM (using M-codeword codebook)

$$\bar{b}_j(k) = \frac{\text{the number of vectors with codebook index } k \text{ in state } j}{\text{the number of vectors in state } j}$$

- For continuous density HMM (M Gaussian mixtures per state)

⇒ cluster the observation vectors within each state j into a set of M clusters

\bar{w}_{jm} = number of vectors classified in cluster m of state j

divided by the number of vectors in state j

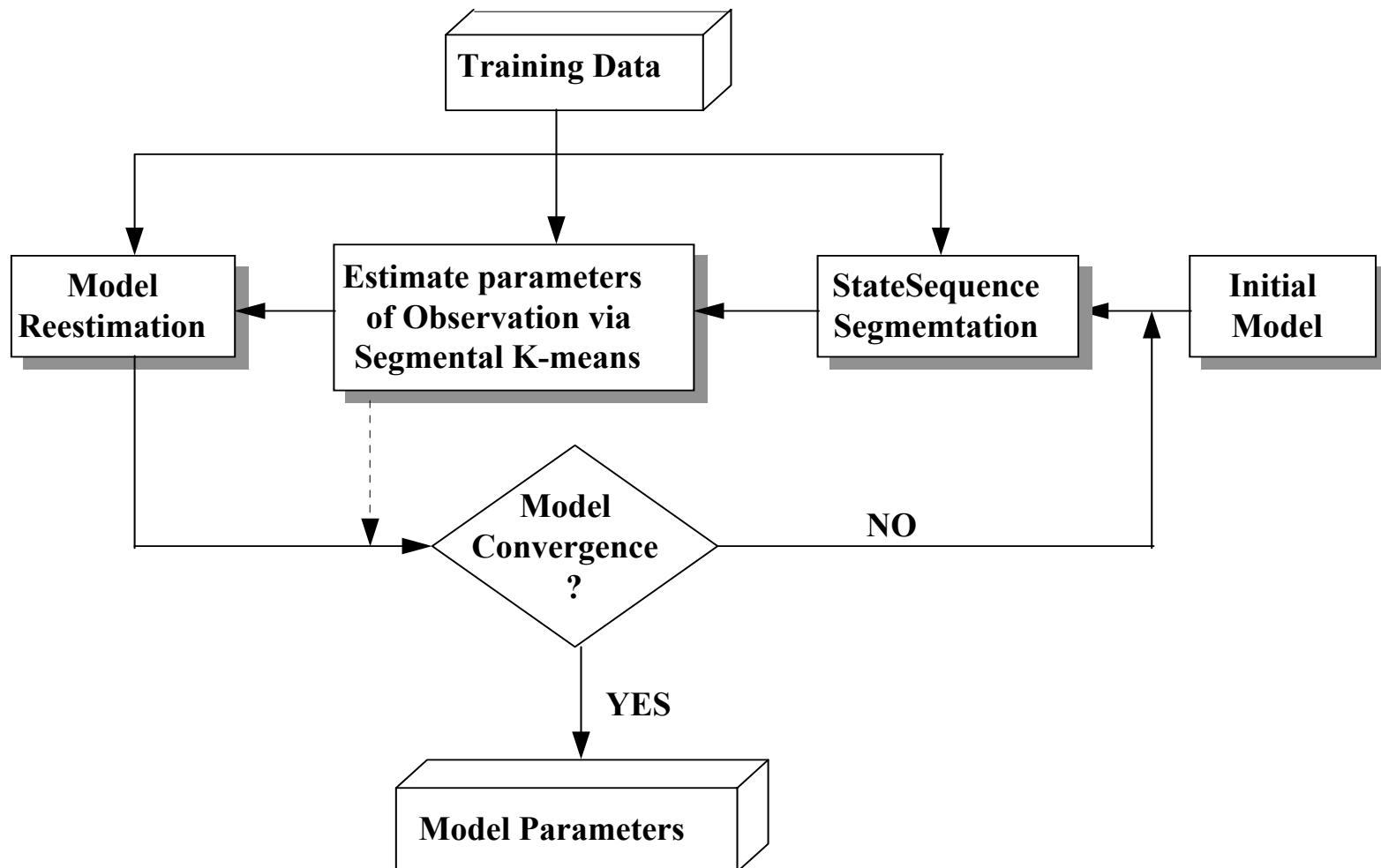
$\bar{\mu}_{jm}$ = sample mean of the vectors classified in cluster m of state j

$\bar{\Sigma}_{jm}$ = sample covariance matrix of the vectors classified in cluster m of state j

- Step 3: Evaluate the model score

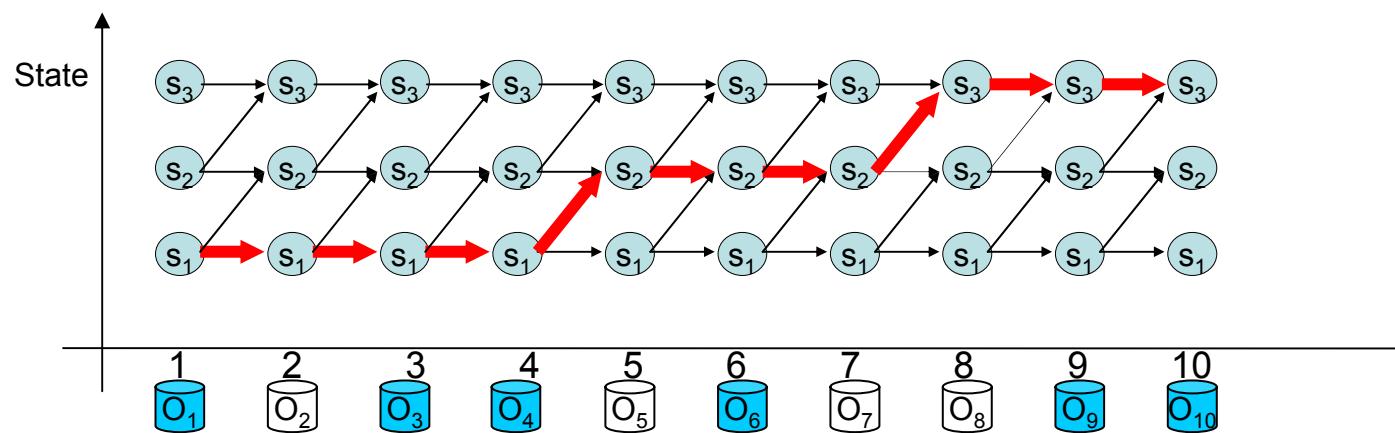
If the difference between the previous and current model scores is greater than a threshold, go back to Step 1, otherwise stop, the initial model is generated

Initialization of HMM (cont.)

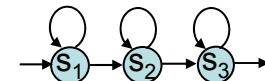
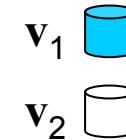


Initialization of HMM (cont.)

- An example for discrete HMM
 - 3 states and 2 codeword

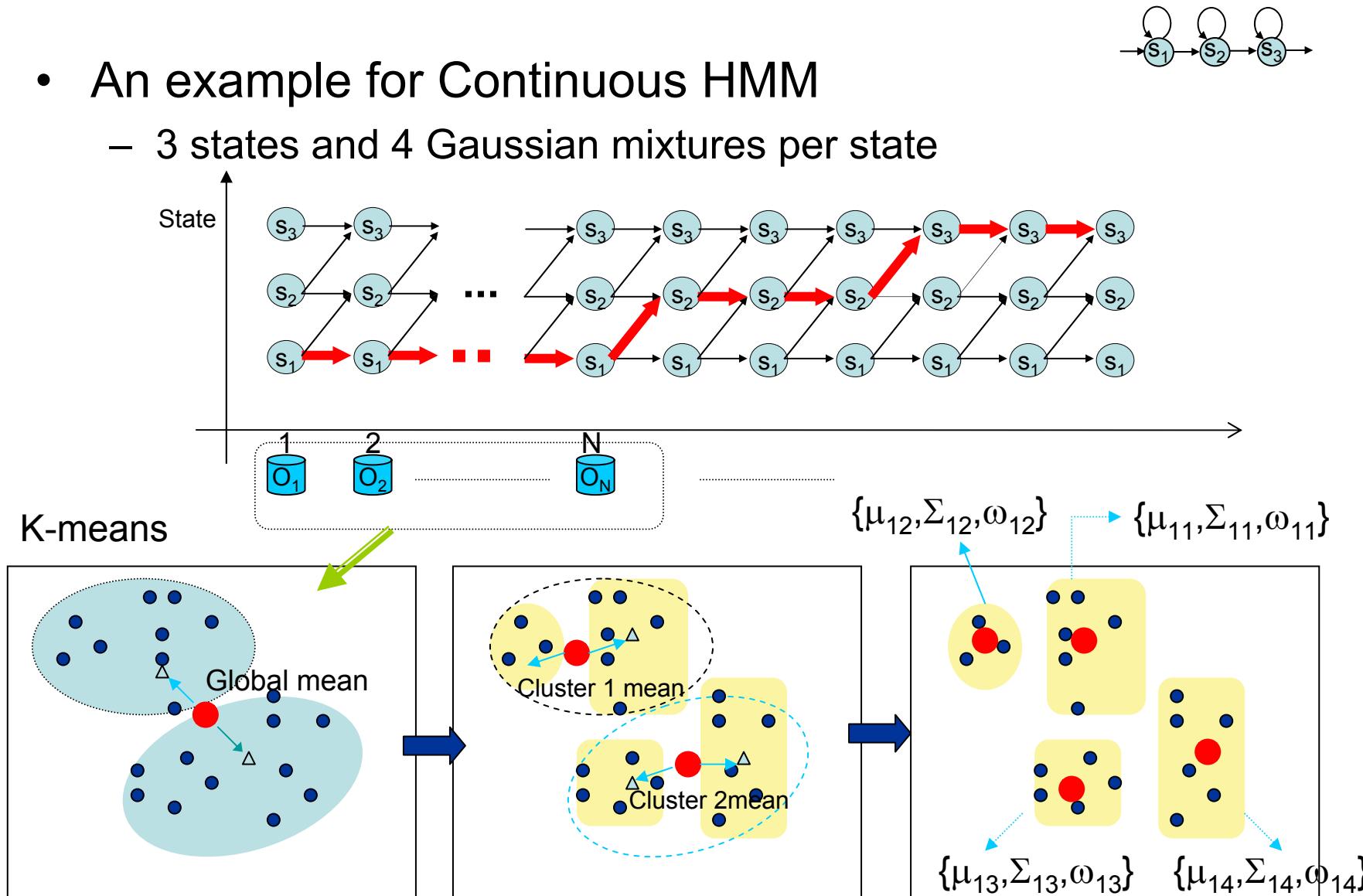


- $b_1(v_1)=3/4$, $b_1(v_2)=1/4$
- $b_2(v_1)=1/3$, $b_2(v_2)=2/3$
- $b_3(v_1)=2/3$, $b_3(v_2)=1/3$



Initialization of HMM (cont.)

- An example for Continuous HMM
 - 3 states and 4 Gaussian mixtures per state



HMM Topology

- Speech is time-evolving non-stationary signal
 - Each HMM state has the ability to capture some quai-stationary segment in the non-stationary speech signal
 - A *left-to-right* topology is a natural candidate to model the speech signal

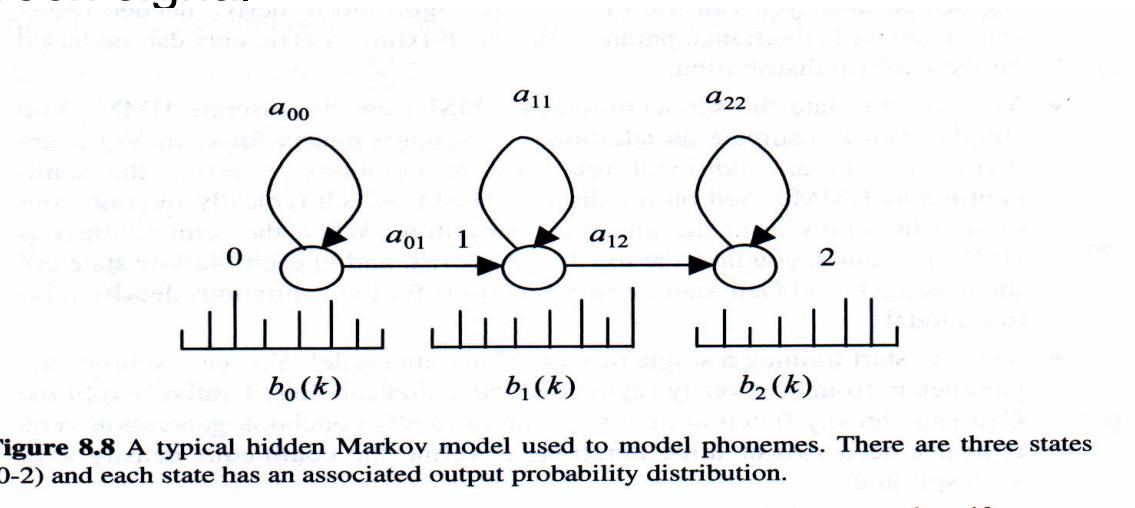


Figure 8.8 A typical hidden Markov model used to model phonemes. There are three states (0-2) and each state has an associated output probability distribution.

- It is general to represent a phone using 3~5 states (English) and a syllable using 6~8 states (Mandarin Chinese)

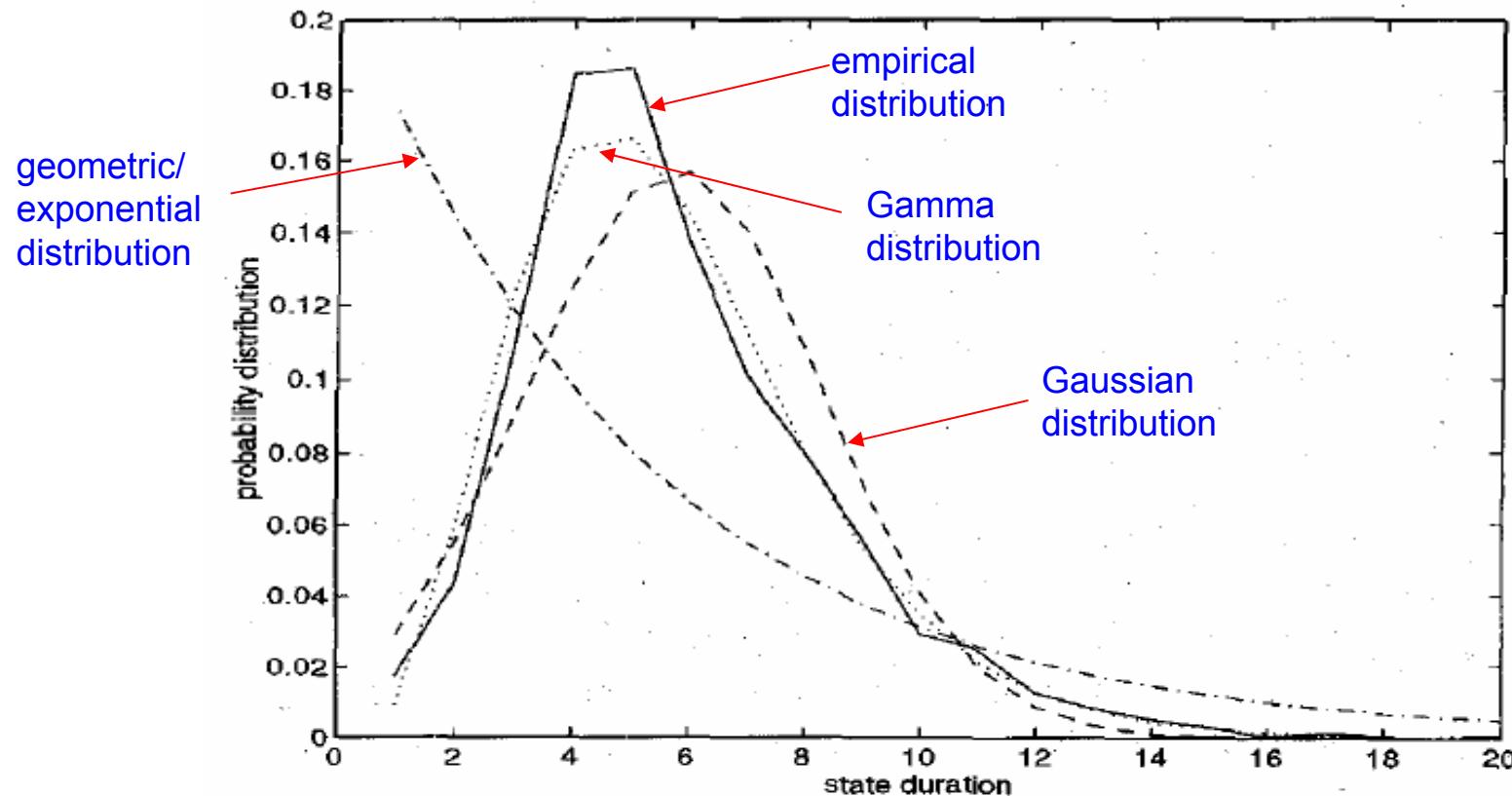
Known Limitations of HMMs

- The assumptions of conventional HMMs in Speech Processing
 - The state duration follows an exponential distribution
 - Don't provide adequate representation of the temporal structure of speech
 - **First-order (Markov) assumption:** the state transition depends only on the origin and destination
 - **Output-independent assumption:** all observation frames are dependent on the state that generated them, not on neighboring observation frames

Researchers have proposed a number of techniques to address these limitations, albeit these solutions have not significantly improved speech recognition accuracy for practical applications.

Known Limitations of HMMs (cont.)

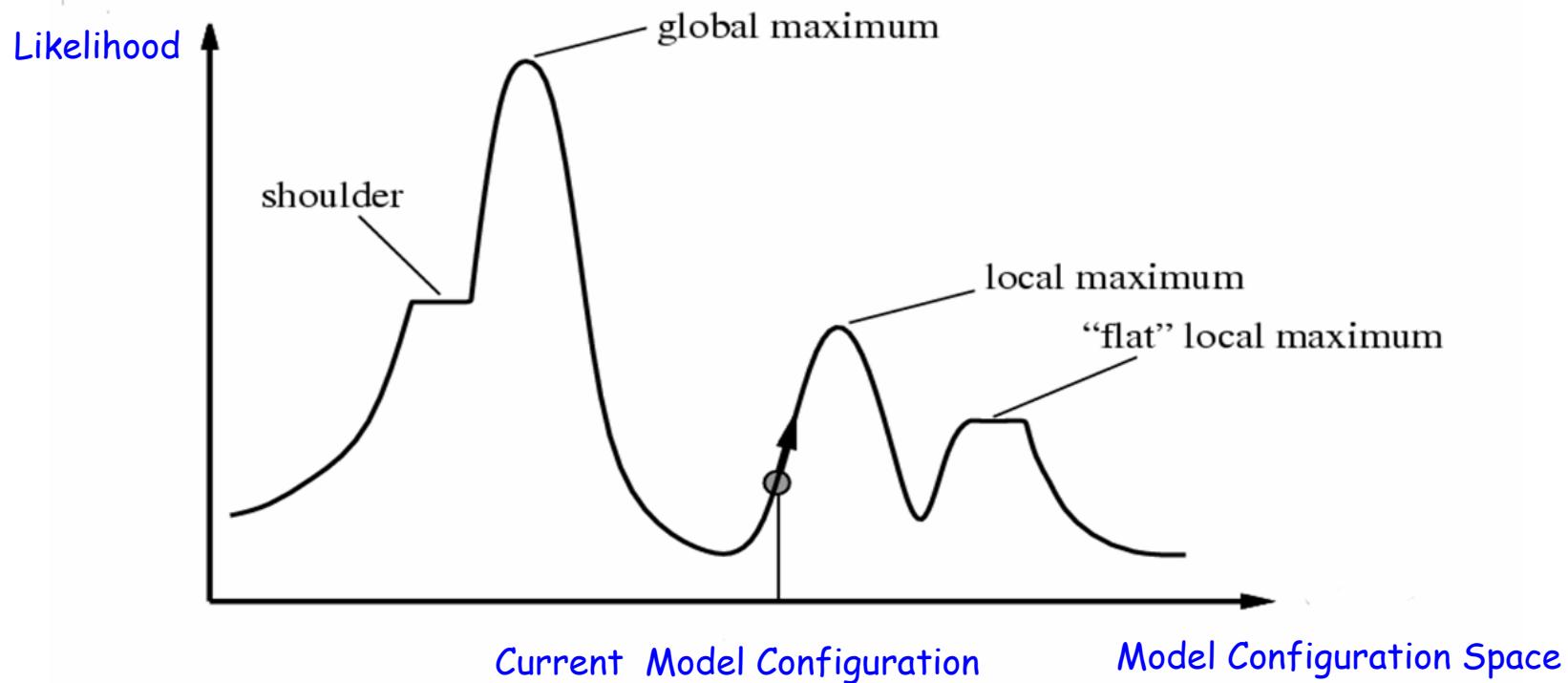
- Duration modeling



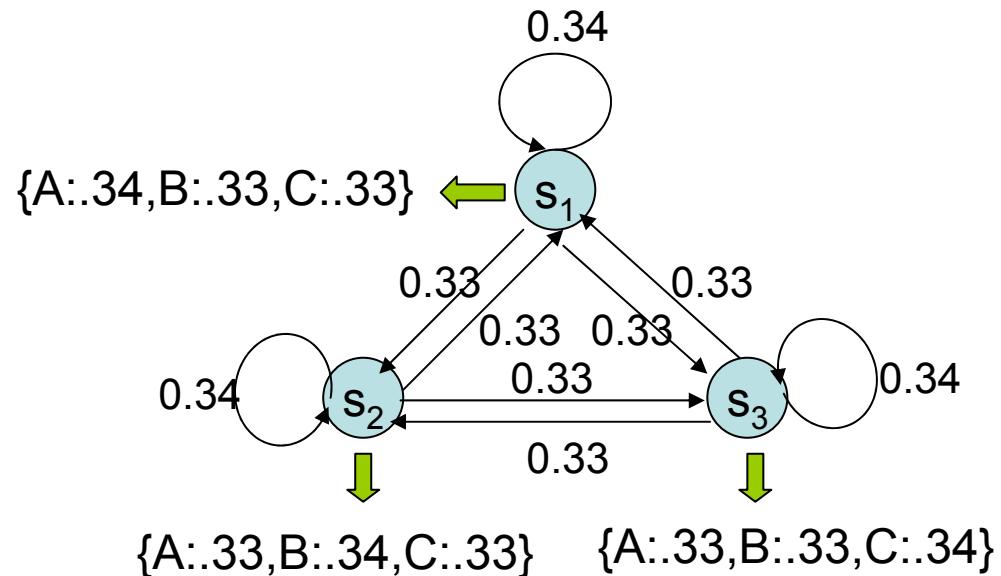
Duration distributions for the seventh state of the word “seven:”
empirical distribution (solid line); Gauss fit (dashed line); gamma fit (dotted line); and (d) geometric fit (dash-dot line).

HMM Limitations (cont.)

- The HMM parameters trained by the *Baum-Welch* algorithm and *EM* algorithm were only locally optimized



Homework-2A



TrainSet 1:

1. ABBCABCABC
2. ABCABC
3. ABCA ABC
4. BBABCAB
5. BCAABCCAB
6. CACCABCA
7. CABCABCABC
8. CABCA
9. CABCA

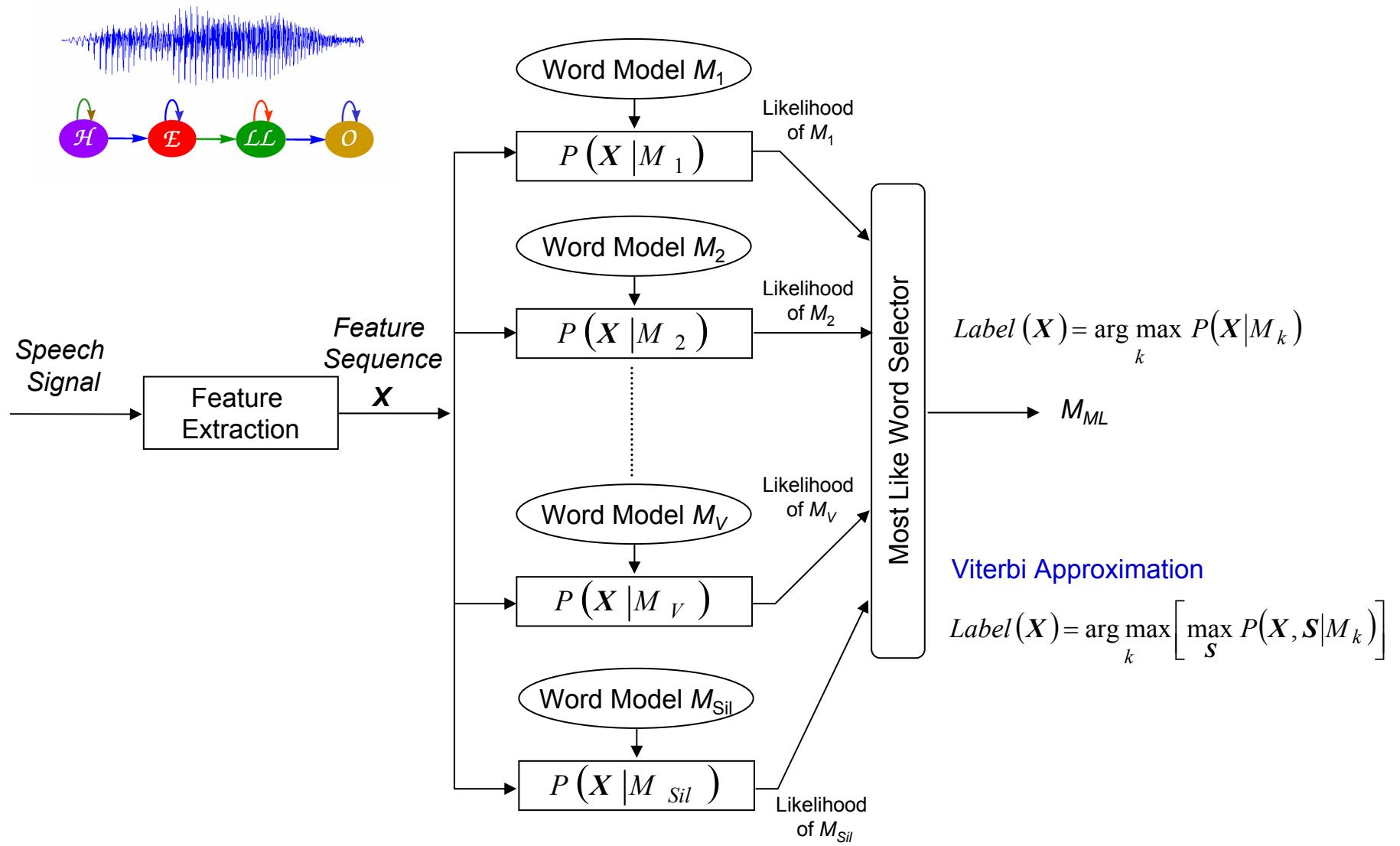
TrainSet 2:

1. BBBCCBC
2. CCBABB
3. AACCBBB
4. BBABBAC
5. CCA ABBAB
6. BBBCCBAA
7. ABBBBABA
8. CCCCC
9. BA AAA

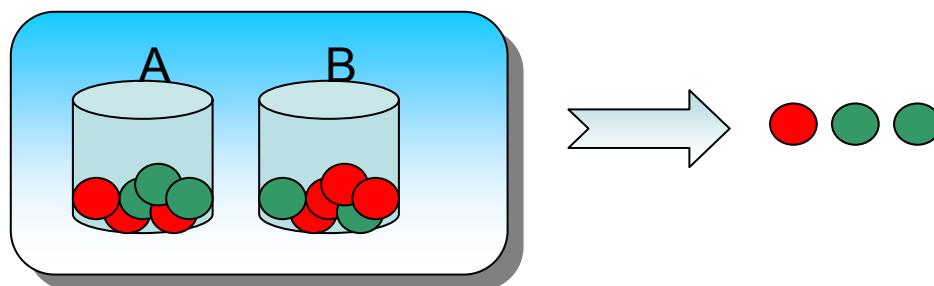
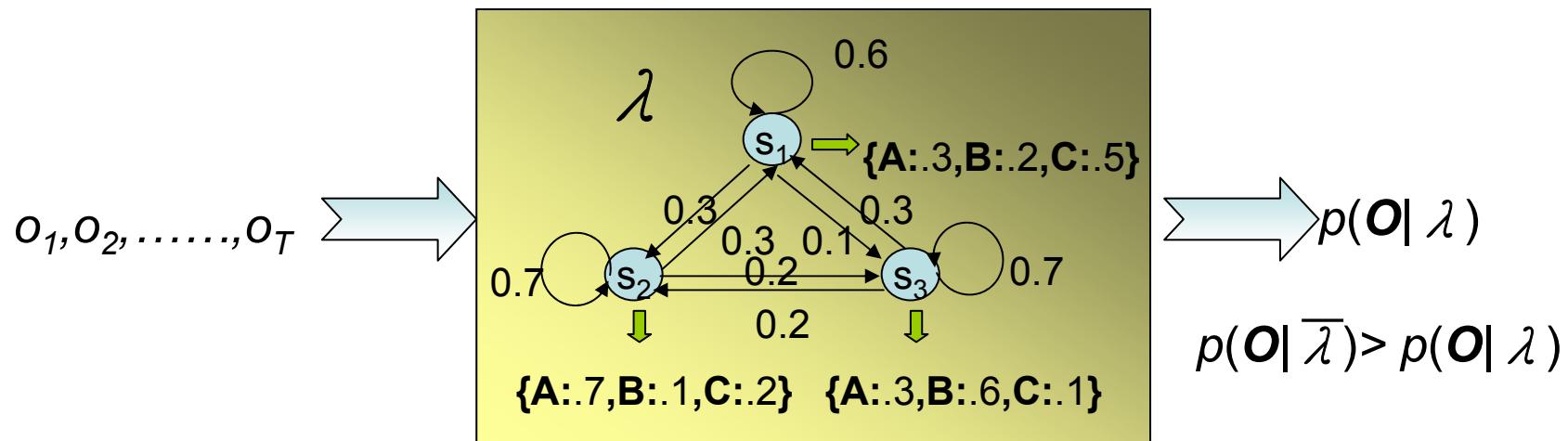
Homework-2A (cont.)

- P1. Please specify the model parameters after the first and 50th iterations of Baum-Welch training
- P2. Please show the recognition results by using the above training sequences as the testing data (The so-called inside testing).
*You have to perform the recognition task with the HMMs trained from the first and 50th iterations of Baum-Welch training, respectively
- P3. Which class do the following testing sequences belong to?
- ABCABCCAB
- AABABCCCCBBB
- P4. What are the results if Observable Markov Models were instead used in P1, P2 and P3?

Isolated Word Recognition



The EM Algorithm



Observed data : O : “ball sequence”
Latent data : S : “bottle sequence”

Parameters to be estimated to maximize $\log P(O|\lambda)$
 $\lambda = \{P(A), P(B), P(B|A), P(A|B), P(R|A), P(G|A), P(R|B), P(G|B)\}$

The EM Algorithm (cont.)

- Introduction of EM (Expectation Maximization):
 - Why EM?
 - Simple optimization algorithms for likelihood function relies on the intermediate variables, called latent (隱藏的) data
In our case here, ***the state sequence is the latent data***
 - Direct access to the data necessary to estimate the parameters is impossible or difficult
In our case here, it is almost impossible to estimate $\{\mathbf{A}, \mathbf{B}, \pi\}$ without consideration of the ***state sequence***
 - Two Major Steps :
 - ***E*** : expectation with respect to the latent data using the current estimate of the parameters and conditioned on the observations $E [\bullet]_{s|\lambda, o}$
 - ***M***: provides a new estimation of the parameters according to Maximum likelihood (ML) or Maximum A Posterior (MAP) Criteria

The EM Algorithm (cont.)

ML and MAP

- Estimation principle based on observations:

$$\boldsymbol{x} = (\boldsymbol{x}_1, \boldsymbol{x}_2, \dots, \boldsymbol{x}_n) \iff X = \{X_1, X_2, \dots, X_n\}$$

- The Maximum Likelihood (ML) Principle**

find the model parameter Φ so that the likelihood $p(\boldsymbol{x}|\Phi)$ is maximum

for example, if $\Phi = \{\mu, \Sigma\}$ is the parameters of a multivariate normal distribution, and \mathbf{X} is i.i.d. (independent, identically distributed), then the ML estimate of $\Phi = \{\mu, \Sigma\}$ is

$$\boldsymbol{\mu}_{ML} = \frac{1}{n} \sum_{i=1}^n \boldsymbol{x}_i, \quad \boldsymbol{\Sigma}_{ML} = \frac{1}{n} \sum_{i=1}^n (\boldsymbol{x}_i - \boldsymbol{\mu}_{ML})(\boldsymbol{x}_i - \boldsymbol{\mu}_{ML})^t$$

- The Maximum A Posteriori (MAP) Principle**

find the model parameter Φ so that the likelihood $p(\Phi|x)$ is maximum

The EM Algorithm (cont.)

- The EM Algorithm is important to HMMs and other learning techniques
 - Discover new model parameters to maximize the log-likelihood of incomplete data $\log P(\mathbf{O}|\lambda)$ by iteratively maximizing the expectation of log-likelihood from complete data $\log P(\mathbf{O}, \mathbf{S}|\lambda)$
- Using scalar random variables to introduce the EM algorithm
 - The observable training data \mathbf{O}
 - We want to maximize $P(\mathbf{O}|\lambda)$, λ is a parameter vector
 - The hidden (unobservable) data \mathbf{S}
 - E.g. the component densities of observable data \mathbf{O} , or the underlying state sequence in HMMs

The EM Algorithm (cont.)

- Assume we have λ and estimate the probability that each s occurred in the generation of o
- Pretend we had in fact observed a complete data pair (o, s) with frequency proportional to the probability $P(o, s | \lambda)$, to compute a new $\bar{\lambda}$, the maximum likelihood estimate of λ
- Does the process converge?
- **Algorithm** unknown model setting

$$P(o, s | \bar{\lambda}) = P(s | o, \bar{\lambda}) P(o | \bar{\lambda})$$

↑ complete data likelihood ↑ incomplete data likelihood

Bayes' rule

- **Log-likelihood expression** and expectation taken over s

$$\log P(o | \bar{\lambda}) = \log P(o, s | \bar{\lambda}) - \log P(s | o, \bar{\lambda})$$

take expectation over s

$$\log P(o | \bar{\lambda}) = \sum_s [P(s | o, \bar{\lambda}) \log P(o | \bar{\lambda})]$$

$$= \sum_s [P(s | o, \bar{\lambda}) \log P(o, s | \bar{\lambda})] - \sum_s [P(s | o, \bar{\lambda}) \log P(s | o, \bar{\lambda})]$$

The EM Algorithm (cont.)

- Algorithm (Cont.)

- We can thus express $\log P(\mathbf{O}|\bar{\lambda})$ as follows

$$\log P(\mathbf{O}|\bar{\lambda})$$

$$= \sum_s [P(\mathbf{S}|\mathbf{O}, \lambda) \log P(\mathbf{O}, \mathbf{S}|\bar{\lambda})] - \sum_s [P(\mathbf{S}|\mathbf{O}, \lambda) \log P(\mathbf{S}|\mathbf{O}, \bar{\lambda})]$$

$$= Q(\lambda, \bar{\lambda}) - H(\lambda, \bar{\lambda})$$

where

$$Q(\lambda, \bar{\lambda}) = \sum_s [P(\mathbf{S}|\mathbf{O}, \lambda) \log P(\mathbf{O}, \mathbf{S}|\bar{\lambda})]$$

$$H(\lambda, \bar{\lambda}) = \sum_s [P(\mathbf{S}|\mathbf{O}, \lambda) \log P(\mathbf{S}|\mathbf{O}, \bar{\lambda})]$$

- We want $\log P(\mathbf{O}|\bar{\lambda}) \geq \log P(\mathbf{O}|\lambda)$

$$\log P(\mathbf{O}|\bar{\lambda}) - \log P(\mathbf{O}|\lambda)$$

$$= [Q(\lambda, \bar{\lambda}) - H(\lambda, \bar{\lambda})] - [Q(\lambda, \lambda) - H(\lambda, \lambda)]$$

$$= Q(\lambda, \bar{\lambda}) - Q(\lambda, \lambda) - H(\lambda, \bar{\lambda}) + H(\lambda, \lambda)$$

The EM Algorithm (cont.)

- $-H(\lambda, \bar{\lambda}) + H(\lambda, \lambda)$ has the following property

$$-H(\lambda, \bar{\lambda}) + H(\lambda, \lambda)$$

$$= -\sum_s \left[P(S|O, \lambda) \log \frac{P(S|O, \bar{\lambda})}{P(S|O, \lambda)} \right]$$

Kullback-Leibler (KL) distance

$$\geq \sum_s \left[P(S|O, \lambda) \left(1 - \frac{P(S|O, \bar{\lambda})}{P(S|O, \lambda)} \right) \right] \quad (\because \log x \leq x - 1)$$

Jensen's inequality

$$= \sum_s [P(S|O, \lambda) - P(S|O, \bar{\lambda})]$$

$$= 0$$

$$\therefore -H(\lambda, \bar{\lambda}) + H(\lambda, \lambda) \geq 0$$

- Therefore, for maximizing $\log P(O|\bar{\lambda})$, we only need to maximize the Q-function (auxiliary function)

$$Q(\lambda, \bar{\lambda}) = \sum_s [P(S|O, \lambda) \log P(O, S|\bar{\lambda})]$$

Expectation of the complete data log likelihood with respect to the latent state sequences

EM Applied to Discrete HMM Training

- Apply EM algorithm to iteratively refine the HMM parameter vector $\lambda = (A, B, \pi)$

– By maximizing the auxiliary function

$$\begin{aligned} Q(\lambda, \bar{\lambda}) &= \sum_s [P(s|o, \lambda) \log P(o, s|\bar{\lambda})] \\ &= \sum_s \left[\frac{P(o, s|\lambda)}{P(o|\lambda)} \log P(o, s|\bar{\lambda}) \right] \end{aligned}$$

– Where $P(o, s|\lambda)$ and $P(o, s|\bar{\lambda})$ can be expressed as

$$P(o, s|\lambda) = \pi_{s_1} \left[\prod_{t=1}^{T-1} a_{s_t s_{t+1}} \right] \left[\prod_{t=1}^T b_{s_t}(o_t) \right]$$

$$\log P(o, s|\lambda) = \log \pi_{s_1} + \sum_{t=1}^{T-1} \log a_{s_t s_{t+1}} + \sum_{t=1}^T \log b_{s_t}(o_t)$$

$$\log P(o, s|\bar{\lambda}) = \log \bar{\pi}_{s_1} + \sum_{t=1}^{T-1} \log \bar{a}_{s_t s_{t+1}} + \sum_{t=1}^T \log \bar{b}_{s_t}(o_t)$$

EM Applied to Discrete HMM Training (cont.)

- Rewrite the auxiliary function as

$$Q(\lambda, \bar{\lambda}) = Q_\pi(\lambda, \bar{\pi}) + Q_a(\lambda, \bar{a}) + Q_b(\lambda, \bar{b})$$

$$Q_\pi(\lambda, \bar{\pi}) = \sum_{\text{all } S} \left[\frac{P(O, S | \lambda)}{P(O | \lambda)} \log \bar{\pi}_{s_1} \right] = \sum_{i=1}^N \left[\frac{P(O, s_1 = i | \lambda)}{P(O | \lambda)} \log \bar{\pi}_i \right]$$

$$Q_a(\lambda, \bar{a}) = \sum_{\text{all } S} \left[\frac{P(O, S | \lambda)}{P(O | \lambda)} \sum_{t=1}^{T-1} \log \bar{a}_{s_t s_{t+1}} \right] = \sum_{i=1}^N \sum_{j=1}^N \sum_{t=1}^{T-1} \left[\frac{P(O, s_t = i, s_{t+1} = j | \lambda)}{P(O | \lambda)} \log \bar{a}_{ij} \right]$$

$$Q_b(\lambda, \bar{b}) = \sum_{\text{all } S} \left[\frac{P(O, S | \lambda)}{P(O | \lambda)} \sum_{t=1}^T \log \bar{b}_{s_t}(k) \right] = \sum_{j=1}^N \sum_k \sum_{t \in o_t = v_k} \left[\frac{P(O, s_t = j | \lambda)}{P(O | \lambda)} \log \bar{b}_j(k) \right]$$

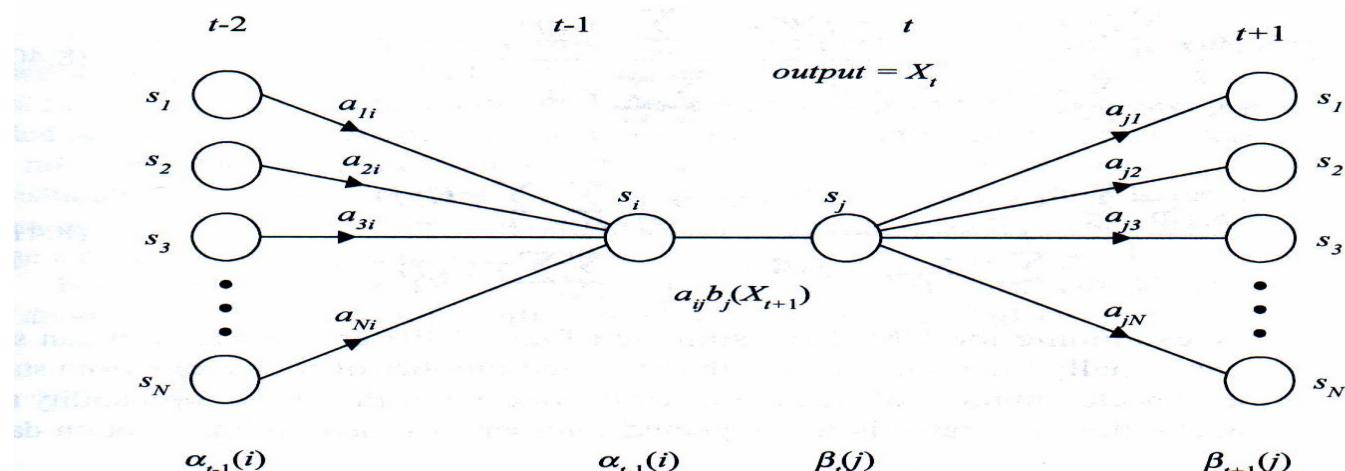


Figure 8.7 Illustration of the operations required for the computation of $\gamma_t(i, j)$, which is the probability of taking the transition from state i to state j at time t .

EM Applied to Discrete HMM Training (cont.)

- The auxiliary function contains three independent terms, π_i , a_{ij} and $b_j(k)$
 - Can be maximized individually
 - All of the same form

$$F(\mathbf{y}) = g(y_1, y_2, \dots, y_N) = \sum_{j=1}^N w_j \log y_j, \text{ where } \sum_{j=1}^N y_j = 1, \text{ and } y_j \geq 0$$

$$F(\mathbf{y}) \text{ has maximum value when : } y_j = \frac{w_j}{\sum_{j=1}^N w_j}$$

EM Applied to Discrete HMM Training (cont.)

- **Proof:** Apply Lagrange Multiplier

By applying Lagrange Multiplier ℓ

Suppose that $F = \sum_{j=1}^N w_j \log y_j = \sum_{j=1}^N w_j \log y_j + \ell \left(\sum_{j=1}^N y_j - 1 \right)$

$$\frac{\partial F}{\partial y_j} = \frac{w_j}{y_j} + \ell = 0 \Rightarrow \ell = -\frac{w_j}{y_j} \quad \forall j$$

Constraint

$$\ell \sum_{j=1}^N y_j = -\sum_{j=1}^N w_j \Rightarrow \ell = -\sum_{j=1}^N w_j$$

$$\therefore y_j = \frac{w_j}{\sum_{j=1}^N w_j}$$

EM Applied to Discrete HMM Training (cont.)

- The new model parameter set $\bar{\lambda} = (\bar{\pi}, \bar{A}, \bar{B})$ can be expressed as:

$$\bar{\pi}_i = \frac{P(\mathbf{o}, s_1 = i | \lambda)}{P(\mathbf{o} | \lambda)} = \gamma_1(i)$$

$$\bar{a}_{ij} = \frac{\sum_{t=1}^{T-1} P(\mathbf{o}, s_t = i, s_{t+1} = j | \lambda)}{\sum_{t=1}^{T-1} P(\mathbf{o}, s_t = i | \lambda)} = \frac{\sum_{t=1}^{T-1} \xi_t(i, j)}{\sum_{t=1}^{T-1} \gamma_t(i)}$$

$$\bar{b}_i(k) = \frac{\sum_{t=1}^T P(\mathbf{o}, s_t = i | \lambda) \text{ s.t. } o_t = v_k}{\sum_{t=1}^T P(\mathbf{o}, s_t = i | \lambda)} = \frac{\sum_{t=1}^T \gamma_t(i) \text{ s.t. } o_t = v_k}{\sum_{t=1}^T \gamma_t(i)}$$

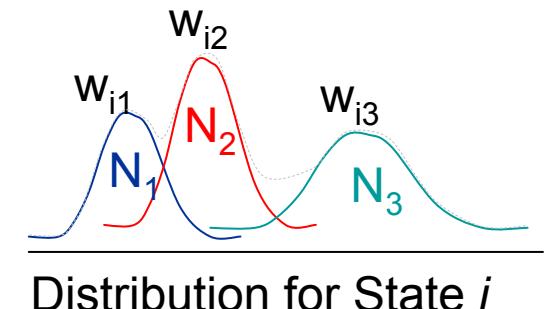
EM Applied to Continuous HMM Training (cont.)

- Continuous HMM: the state observation does not come from a finite set, but from a continuous space
 - The difference between the discrete and continuous HMM lies in a different form of state output probability
 - Discrete HMM requires the quantization procedure to map observation vectors from the continuous space to the discrete space
- Continuous Mixture HMM
 - The state observation distribution of HMM is modeled by multivariate Gaussian mixture density functions (M mixtures)

$$b_j(\mathbf{o}) = \sum_{k=1}^M c_{jk} b_{jk}(\mathbf{o})$$

$$= \sum_{k=1}^M c_{jk} N(\mathbf{o}; \boldsymbol{\mu}_{jk}, \boldsymbol{\Sigma}_{jk}) = \sum_{k=1}^M c_{jk} \left(\frac{1}{(\sqrt{2\pi})^L |\boldsymbol{\Sigma}_{jk}|^{1/2}} \exp\left(-\frac{1}{2} (\mathbf{o} - \boldsymbol{\mu}_{jk})^\top \boldsymbol{\Sigma}_{jk}^{-1} (\mathbf{o} - \boldsymbol{\mu}_{jk})\right) \right)$$

$$\sum_{k=1}^M c_{jk} = 1$$



EM Applied to Continuous HMM Training (cont.)

- Express $b_j(\mathbf{o})$ with respect to each single mixture component $b_{jk}(\mathbf{o})$

$$P(\mathbf{O}, \mathbf{S} | \lambda) = \pi_{s_1} \left\{ \prod_{t=1}^{T-1} a_{s_t s_{t+1}} \right\} \left\{ \prod_{t=1}^T b_{s_t}(\mathbf{o}_t) \right\}$$

Note:

$$\prod_{t=1}^T \left(\sum_{k_t=1}^M a_{tk_t} \right)$$

 $= (a_{11} + a_{12} + \dots + a_{1M})(a_{21} + a_{22} + \dots + a_{2M}) \dots (a_{T1} + a_{T2} + \dots + a_{TM})$
 $= \sum_{k_1=1}^M \sum_{k_2=1}^M \dots \sum_{k_T=1}^M \prod_{t=1}^T a_{tk_t}$



$$\downarrow \quad = \pi_{s_1} \left\{ \prod_{t=1}^{T-1} a_{s_t s_{t+1}} \right\} \left\{ \sum_{k_1=1}^M \sum_{k_2=1}^M \dots \sum_{k_T=1}^M \prod_{t=1}^T [c_{s_t k_t} b_{s_t k_t}(\mathbf{o}_t)] \right\}$$

$$P(\mathbf{O}, \mathbf{S}, \mathbf{K} | \lambda) = \pi_{s_1} \left\{ \prod_{t=1}^{T-1} a_{s_t s_{t+1}} \right\} \left\{ \prod_{t=1}^T [c_{s_t k_t} b_{s_t k_t}(\mathbf{o}_t)] \right\}$$

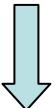
\mathbf{K} : one of the possible mixture component sequence
along with the state sequence \mathbf{S}

$$P(\mathbf{O} | \lambda) = \sum_{\mathbf{S}} \sum_{\mathbf{K}} P(\mathbf{O}, \mathbf{S}, \mathbf{K} | \lambda)$$

EM Applied to Continuous HMM Training (cont.)

- Therefore, an auxiliary function for the EM algorithm can be written as:

$$\begin{aligned} Q(\lambda, \bar{\lambda}) &= \sum_s \sum_K \left[P(S, K | O, \lambda) \log P(O, S, K | \bar{\lambda}) \right] \\ &= \sum_s \sum_K \left[\frac{P(O, S, K | \lambda)}{P(O | \lambda)} \log P(O, S, K | \bar{\lambda}) \right] \end{aligned}$$

$$\log P(O, S, K | \bar{\lambda}) = \log \bar{\pi}_{s_1} + \sum_{t=1}^{T-1} \log \bar{a}_{s_t s_{t+1}} + \sum_{t=1}^T \log \bar{b}_{s_t k_t}(o_t) + \sum_{t=1}^T \log \bar{c}_{s_t k_t}$$


$$Q(\lambda, \bar{\lambda}) = Q_\pi(\lambda, \bar{\pi}) + Q_a(\lambda, \bar{a}) + Q_b(\lambda, \bar{b}) + Q_c(\lambda, \bar{c})$$

initial
probabilities

state transition
probabilities

Gaussian
density
functions

mixture
components

EM Applied to Continuous HMM Training (cont.)

- The only difference we have when compared with Discrete HMM training

$$Q_b(\lambda, \bar{b}) = \sum_{t=1}^T \left\{ \left[\sum_{j=1}^N \sum_{k=1}^M P(s_t = j, k_t = k | \mathbf{o}, \lambda) \right] \log \bar{b}_{jk}(\mathbf{o}_t) \right\}$$

$$Q_c(\lambda, \bar{c}) = \sum_{t=1}^T \left\{ \left[\sum_{j=1}^N \sum_{k=1}^M P(s_t = j, k_t = k | \mathbf{o}, \lambda) \right] \log \bar{c}_{jk}(\mathbf{o}_t) \right\}$$

EM Applied to Continuous HMM Training (cont.)

Let $\gamma_t(j, k) = \sum_{k=1}^M P(s_t = j, k_t = k | \mathbf{o}_t, \lambda)$

$$\bar{b}_{jk}(\mathbf{o}_t) = N(\mathbf{o}_t; \bar{\boldsymbol{\mu}}_{jk}, \bar{\Sigma}_{jk}) = \frac{1}{(2\pi)^{L/2} |\bar{\Sigma}_{jk}|^{1/2}} \exp \left(-\frac{1}{2} (\mathbf{o}_t - \bar{\boldsymbol{\mu}}_{jk})^\top \bar{\Sigma}_{jk}^{-1} (\mathbf{o}_t - \bar{\boldsymbol{\mu}}_{jk}) \right)$$

$$\log \bar{b}_{jk}(\mathbf{o}_t) = -\frac{L}{2} \cdot \log (2\pi) + \frac{1}{2} \cdot \log |\bar{\Sigma}_{jk}^{-1}| - \left(\frac{1}{2} (\mathbf{o}_t - \bar{\boldsymbol{\mu}}_{jk})^\top \bar{\Sigma}_{jk}^{-1} (\mathbf{o}_t - \bar{\boldsymbol{\mu}}_{jk}) \right)$$

$$\frac{\partial \log \bar{b}_{jk}(\mathbf{o}_t)}{\partial \bar{\boldsymbol{\mu}}_{jk}} = \bar{\Sigma}_{jk}^{-1} (\mathbf{o}_t - \bar{\boldsymbol{\mu}}_{jk})$$

$$\frac{\partial Q_b(\lambda, \bar{\mathbf{b}})}{\partial \bar{\boldsymbol{\mu}}_{jk}} = \frac{\partial \sum_{t=1}^T \left\{ \left[\sum_{j=1}^N \sum_{k=1}^M \gamma_t(j, k) \log \bar{b}_{jk}(\mathbf{o}_t) \right] \right\}}{\partial \bar{\boldsymbol{\mu}}_{jk}}$$

$$\Rightarrow \sum_{t=1}^T \left\{ \gamma_t(j, k) \bar{\Sigma}_{jk}^{-1} (\mathbf{o}_t - \bar{\boldsymbol{\mu}}_{jk}) \right\} = 0$$

$$\Rightarrow \bar{\boldsymbol{\mu}}_{jk} = \frac{\sum_{t=1}^T [\gamma_t(j, k) \cdot \mathbf{o}_t]}{\sum_{t=1}^T \gamma_t(j, k)}$$

$$\frac{d(\mathbf{x}^T \mathbf{C} \mathbf{x})}{d\mathbf{x}} = (\mathbf{C} + \mathbf{C}^T) \mathbf{x}$$

and Σ_{jk}^{-1} is symmetric here

EM Applied to Continuous HMM Training (cont.)

$$\log \bar{b}_{jk}(\mathbf{o}_t) = -\frac{1}{2} \cdot \log(2\pi) - \frac{1}{2} \cdot \log |\bar{\Sigma}_{jk}| - \left(\frac{1}{2} (\mathbf{o}_t - \bar{\mu}_{jk})^\top \bar{\Sigma}_{jk}^{-1} (\mathbf{o}_t - \bar{\mu}_{jk}) \right)$$

$$\frac{\partial \log \bar{b}_{jk}(\mathbf{o}_t)}{\partial (\bar{\Sigma}_{jk})} = - \left[\frac{1}{2} \cdot \cancel{|\bar{\Sigma}_{jk}|^{-1}} \cdot \cancel{|\bar{\Sigma}_{jk}|} \cdot \bar{\Sigma}_{jk}^{-1} - \left(\bar{\Sigma}_{jk}^{-1} \frac{1}{2} (\mathbf{o}_t - \bar{\mu}_{jk}) (\mathbf{o}_t - \bar{\mu}_{jk})^\top \bar{\Sigma}_{jk}^{-1} \right) \right]$$

$$= -\frac{1}{2} \cdot \left[\bar{\Sigma}_{jk}^{-1} - \bar{\Sigma}_{jk}^{-1} (\mathbf{o}_t - \bar{\mu}_{jk}) (\mathbf{o}_t - \bar{\mu}_{jk})^\top \bar{\Sigma}_{jk}^{-1} \right]$$

$$\frac{\partial Q_b(\lambda, \bar{b})}{\partial (\bar{\Sigma}_{jk})} = \frac{\partial \sum_{t=1}^T \left\{ \left[\sum_{j=1}^N \sum_{k=1}^M \gamma_t(j,k) \log \bar{b}_{jk}(\mathbf{o}_t) \right] \right\}}{\partial (\bar{\Sigma}_{jk}^{-1})}$$

$$\frac{d(\mathbf{a}^T \mathbf{X}^{-1} \mathbf{b})}{d\mathbf{X}} = -\mathbf{X}^T \mathbf{a} \mathbf{b}^T \mathbf{X}^T$$

$$\frac{d[\det(\mathbf{X})]}{d\mathbf{X}} = \det(\mathbf{X}) \cdot \mathbf{X}^{-T}$$

and Σ_{jk} is symmetric here

$$\Rightarrow \sum_{t=1}^T \left\{ \gamma_t(j,k) \left(-\frac{1}{2} \right) \cdot \left[\bar{\Sigma}_{jk}^{-1} - \bar{\Sigma}_{jk}^{-1} (\mathbf{o}_t - \bar{\mu}_{jk}) (\mathbf{o}_t - \bar{\mu}_{jk})^\top \bar{\Sigma}_{jk}^{-1} \right] \right\} = 0$$

$$\Rightarrow \sum_{t=1}^T \gamma_t(j,k) \bar{\Sigma}_{jk}^{-1} = \sum_{t=1}^T \gamma_t(j,k) \bar{\Sigma}_{jk}^{-1} (\mathbf{o}_t - \bar{\mu}_{jk}) (\mathbf{o}_t - \bar{\mu}_{jk})^\top \bar{\Sigma}_{jk}^{-1}$$

$$\Rightarrow \sum_{t=1}^T \gamma_t(j,k) \bar{\Sigma}_{jk} \bar{\Sigma}_{jk}^{-1} \bar{\Sigma}_{jk} = \sum_{t=1}^T \gamma_t(j,k) \bar{\Sigma}_{jk} \bar{\Sigma}_{jk}^{-1} (\mathbf{o}_t - \bar{\mu}_{jk}) (\mathbf{o}_t - \bar{\mu}_{jk})^\top \bar{\Sigma}_{jk}^{-1} \bar{\Sigma}_{jk}$$

$$\Rightarrow \bar{\Sigma}_{jk} = \frac{\sum_{t=1}^T [\gamma_t(j,k) \cdot (\mathbf{o}_t - \bar{\mu}_{jk}) (\mathbf{o}_t - \bar{\mu}_{jk})^\top]}{\sum_{t=1}^T \gamma_t(j,k)}$$

EM Applied to Continuous HMM Training (cont.)

- The new model parameter set for each mixture component and mixture weight can be expressed as:

$$\bar{\boldsymbol{\mu}}_{jk} = \frac{\sum_{t=1}^T \left[\frac{P(\mathbf{o}, s_t = j, k_t = k | \lambda)}{P(\mathbf{o} | \lambda)} \mathbf{o}_t \right]}{\sum_{t=1}^T \frac{P(\mathbf{o}, s_t = j, k_t = k | \lambda)}{P(\mathbf{o} | \lambda)}} = \frac{\sum_{t=1}^T \gamma_t(j, k) \mathbf{o}_t}{\sum_{t=1}^T \gamma_t(j, k)}$$

$$\bar{\Sigma}_{jk} = \frac{\sum_{t=1}^T \left[\frac{P(\mathbf{o}, s_t = j, k_t = k | \lambda)}{P(\mathbf{o} | \lambda)} (\mathbf{o}_t - \bar{\boldsymbol{\mu}}_{jk})(\mathbf{o}_t - \bar{\boldsymbol{\mu}}_{jk})^T \right]}{\sum_{t=1}^T \frac{P(\mathbf{o}, s_t = j, k_t = k | \lambda)}{P(\mathbf{o} | \lambda)}} = \frac{\sum_{t=1}^T \gamma_t(j, k) (\mathbf{o}_t - \bar{\boldsymbol{\mu}}_{jk})(\mathbf{o}_t - \bar{\boldsymbol{\mu}}_{jk})^T}{\sum_{t=1}^T \gamma_t(j, k)}$$

$$\bar{c}_{jk} = \frac{\sum_{t=1}^T \gamma_t(j, k)}{\sum_{t=1}^T \sum_{k=1}^M \gamma_t(j, k)}$$

Measures of ASR Performance

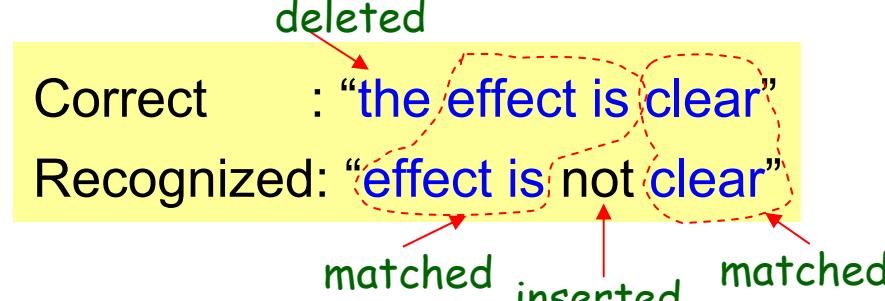
- Evaluating the performance of automatic speech recognition (ASR) systems is critical, and the **Word Recognition Error Rate (WER)** is one of the most important measures
- There are typically three types of word recognition errors
 - **Substitution**
 - An incorrect word was substituted for the correct word
 - **Deletion**
 - A correct word was omitted in the recognized sentence
 - **Insertion**
 - An extra word was added in the recognized sentence
- How to determine the minimum error rate?

Measures of ASR Performance (cont.)

- Calculate the WER by aligning the correct word string against the recognized word string

- A maximum substring matching problem
 - Can be handled by dynamic programming

- Example:



- Error analysis: one deletion and one insertion
 - Measures: word error rate (WER), word correction rate (WCR), word accuracy rate (WAR)

Might be higher than 100%

$$\text{Word Error Rate} = 100\% \frac{\text{Sub.} + \text{Del.} + \text{Ins. words}}{\text{No. of words in the correct sentence}} = \frac{2}{4} = 50\%$$

WER+
WAR
=100%

$$\text{Word Correction Rate} = 100\% \frac{\text{Matched words}}{\text{No. of words in the correct sentence}} = \frac{3}{4} = 75\%$$

$$\text{Word Accuracy Rate} = 100\% \frac{\text{Matched - Ins. words}}{\text{No. of words in the correct sentence}} = \frac{3 - 1}{4} = 50\%$$

Might be negative
SP 2004 - Berlin Chen 85

Measures of ASR Performance (cont.)

- A Dynamic Programming Algorithm (Textbook)

ALGORITHM 9.1: ALGORITHM TO MEASURE THE WORD ERROR RATE

Step 1: Initialization $R[0,0] = 0$ $R[i,j] = \infty$ if ($i < 0$) or ($j < 0$) $B[0,0] = 0$

Step 2: Iteration

for $i = 1, \dots, n$ { //denotes for the word length of the correct/reference sentence

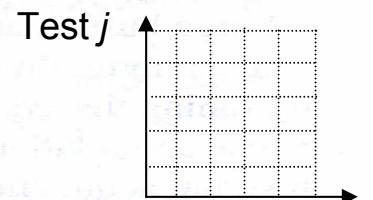
 for $j = 1, \dots, m$ //denotes for the word length of the recognized/test sentence

minimum word
error alignment
at the a grid $[i,j]$

$$R[i, j] = \min \begin{bmatrix} R[i-1, j] + 1 \text{ (deletion)} \\ R[i-1, j-1] \text{ (match) / hit} \\ R[i-1, j-1] + 1 \text{ (substitution)} \\ R[i, j-1] + 1 \text{ (insertion)} \end{bmatrix}$$

kinds of
alignment

$$B[i, j] = \begin{cases} 1 & \text{if deletion} \\ 2 & \text{if insertion} \\ 3 & \text{if match / hit} \\ 4 & \text{if substitution} \end{cases}$$



Step 3: Backtracking and termination

$$\text{word error rate} = 100\% \times \frac{R(n,m)}{n}$$

$$\text{optimal backward path} = (s_1, s_2, \dots, 0)$$

$$\text{where } s_1 = B[n,m], s_t = \begin{cases} B[i-1, j] \text{ if } s_{t-1} = 1 \\ B[i, j-1] \text{ if } s_{t-1} = 2 \\ B[i-1, j-1] \text{ if } s_{t-1} = 3 \text{ or } 4 \end{cases} \text{ for } t = 2, \dots \text{ until } s_t = 0$$

Measures of ASR Performance (cont.)

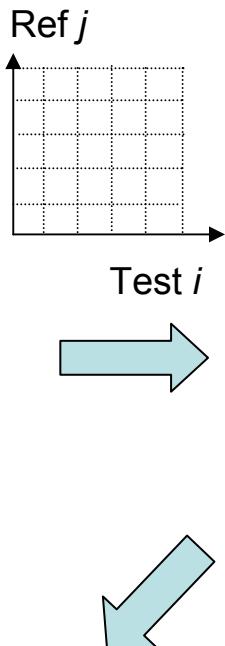
- Algorithm (by Berlin Chen)

Step 1: Initialization :

```

G[0][0] = 0;
for i = 1,..., n { //test
    G[i][0] = G[i-1][0] + 1;
    B[i][0] = 1; //Insertion
} (Horizontal Direction)
for j = 1,..., m { //reference
    G[0][j] = G[0][j-1] + 1;
    B[0][j] = 2; // Deletion
} (Vertical Direction)

```



Step 2 : Iteration :

```

for i = 1,..., n { //test
    for j = 1,..., m { //reference
        G[i][j] = min [
            G[i-1][j] + 1 (Insertion)
            G[i][j-1] + 1 (Deletion)
            G[i-1][j-1] + 1 (if LR[j] != LT[i], Substitution)
            G[i-1][j-1] (if LR[j] = LT[i], Match)
        ]
        B[i][j] = {
            1; //Insertion, (Horizontal Direction)
            2; //Deletion , (Vertical Direction)
            3; //Substitution (Diagonal Direction)
            4; //match (Diagonal Direction)
        }
    } //for j, reference
} //for i, test

```

Step 3 : Measure and Backtrace :

$$\text{Word Error Rate} = 100\% \times \frac{G[n][m]}{m}$$

$$\text{Word Accuracy Rate} = 100\% - \text{Word Error Rate}$$

Optimal backtrace path = $(B[n][m] \rightarrow \dots \rightarrow B[0][0])$

if $B[i][j] = 1$ print " $LT[i]$ "; //Insertion, then go left

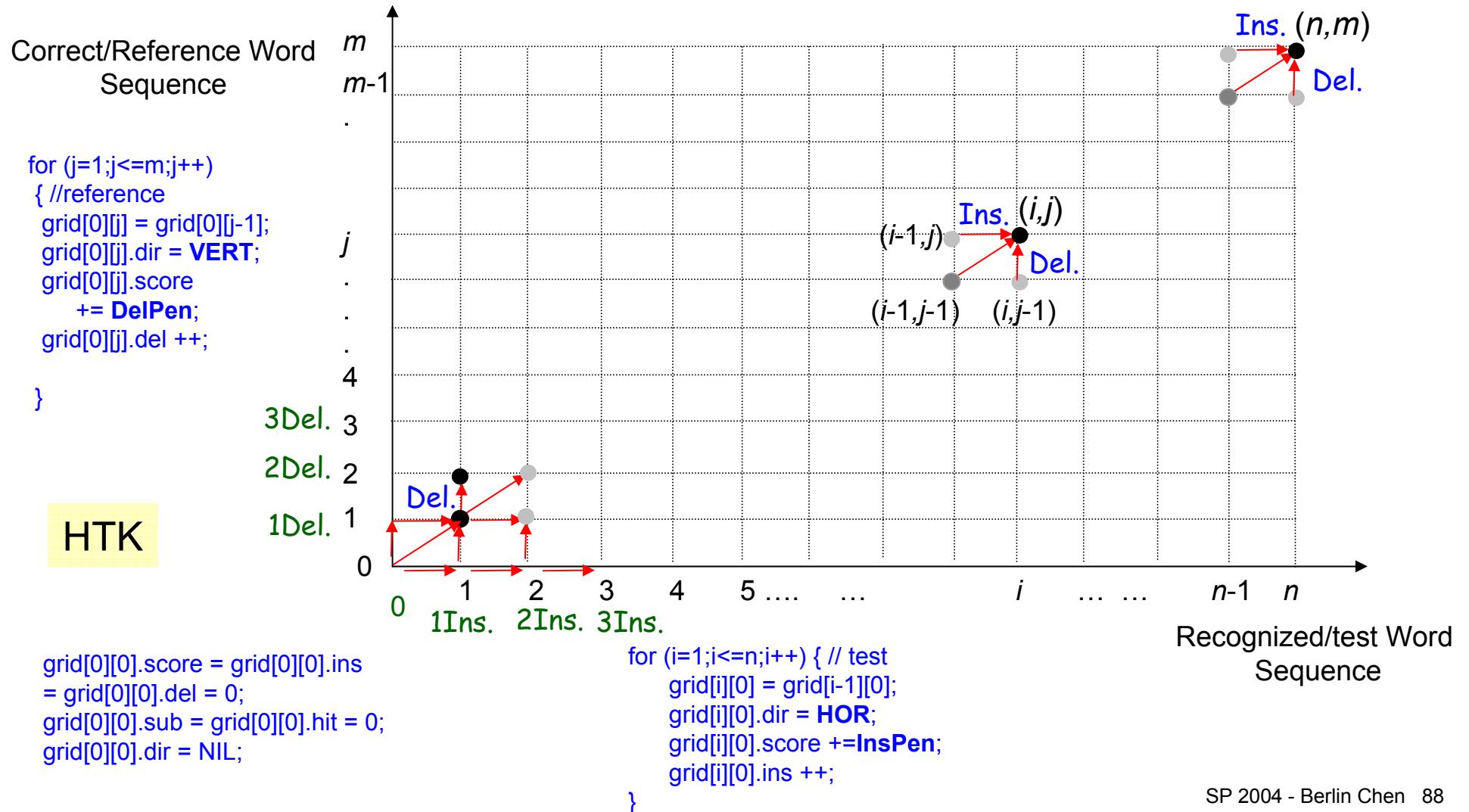
else if $B[i][j] = 2$ print " $LR[j]$ "; //Deletion, then go down

else print " $LR[j] LT[i]$ "; //Hit/Match or Substitution, then go down diagonally

Note: the penalties for substitution, deletion and insertion errors are all set to be 1 here

Measures of ASR Performance (cont.)

- A Dynamic Programming Algorithm
 - Initialization



Measures of ASR Performance (cont.)

- Example 2

Note: the penalties for substitution, deletion and insertion errors are all set to be 1 here

(Ins,Del,Sub,Hit)

Alignment 1: WER= 80%

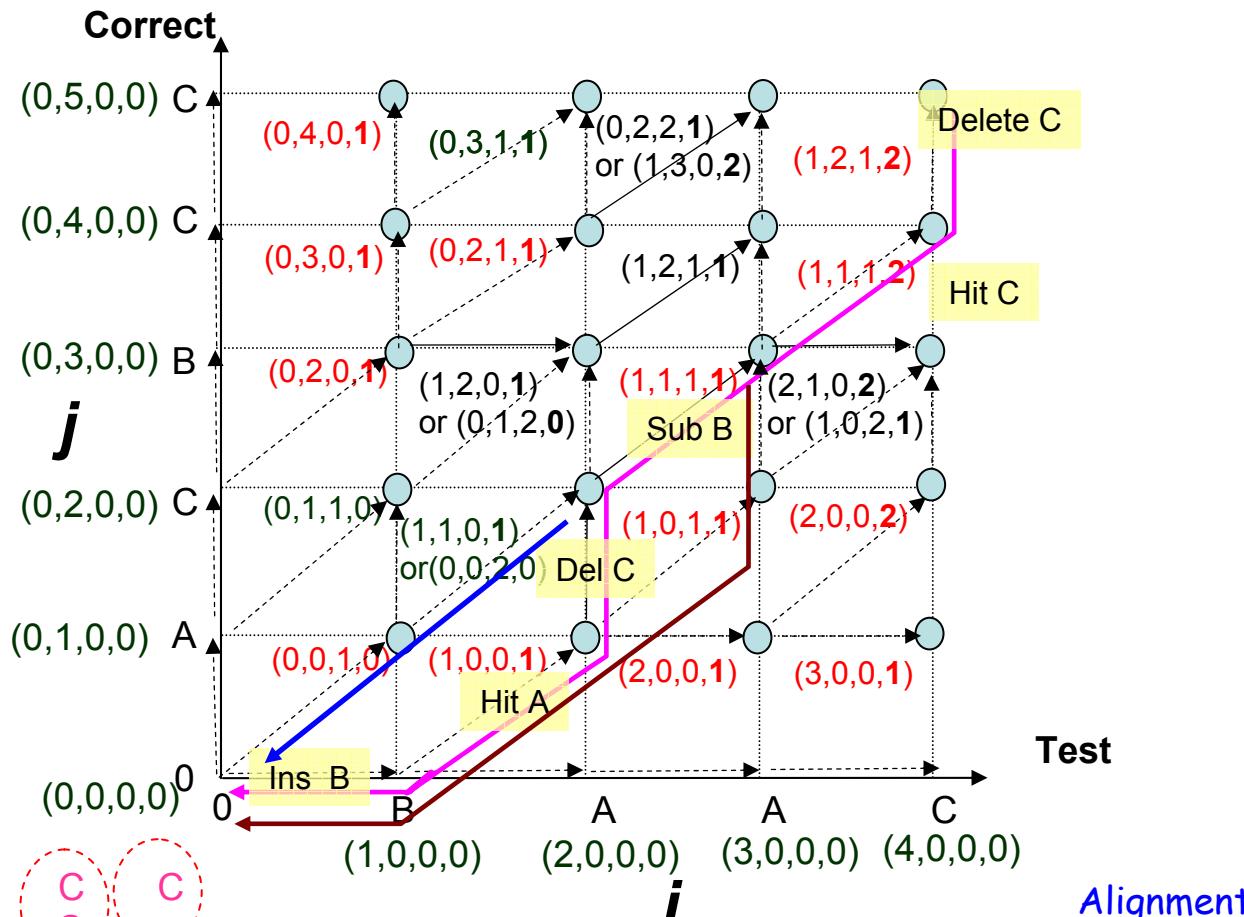
Correct:
Test:

Ins B Hit A Del C Sub B Hit C Del C

Correct:
Test:

Sub A Sub C Sub B Hit C Del C

Alignment 2:
WER=80%



Alignment 3:
WER=80%

Correct:
Test:

Ins B Hit A Sub C Del B Hit C Del C

Measures of ASR Performance (cont.)

- Two common settings of different penalties for substitution, deletion, and insertion errors

```
/* HTK error penalties */  
subPen = 10;  
delPen = 7;  
insPen = 7;  
  
/* NIST error penalties */  
subPenNIST = 4;  
delPenNIST = 3;  
insPenNIST = 3;
```

Homework-2B

- Measures of ASR Performance

Reference

100000 100000 桃
100000 100000 芝
100000 100000 鮑
100000 100000 風
100000 100000 重
100000 100000 創
100000 100000 花
100000 100000 蓮
100000 100000 光
100000 100000 復
100000 100000 鄉
100000 100000 大
100000 100000 興
100000 100000 村
100000 100000 死
100000 100000 傷
100000 100000 慘
100000 100000 重
100000 100000 感
100000 100000 觸
100000 100000 最
100000 100000 多

ASR Output

100000 100000 桃
100000 100000 芝
100000 100000 鮑
100000 100000 風
100000 100000 重
100000 100000 創
100000 100000 花
100000 100000 蓮
100000 100000 光
100000 100000 復
100000 100000 鄉
100000 100000 打
100000 100000 新
100000 100000 村
100000 100000 次
100000 100000 傷
100000 100000 殘
100000 100000 周
100000 100000 感
100000 100000 觸
100000 100000 最
100000 100000 多

.....

.....

Homework-2B (count.)

- 506 BN stories of ASR outputs
 - Report the CER (character error rate) of the first one, 100, 200, and 506 stories
 - The result should show the number of substitution, deletion and insertion errors

----- Overall Results -----

SENT: %Correct=0.00 [H=0, S=1, N=1]
WORD: %Corr=81.52, Acc=81.52 [H=75, D=4, S=13, I=0, N=92]

=====

----- Overall Results -----

SENT: %Correct=0.00 [H=0, S=100, N=100]
WORD: %Corr=87.66, Acc=86.83 [H=10832, D=177, S=1348, I=102, N=12357]

=====

----- Overall Results -----

SENT: %Correct=0.00 [H=0, S=200, N=200]
WORD: %Corr=87.91, Acc=87.18 [H=22657, D=293, S=2824, I=186, N=25774]

=====

----- Overall Results -----

SENT: %Correct=0.00 [H=0, S=506, N=506]
WORD: %Corr=86.83, Acc=86.06 [H=57144, D=829, S=7839, I=504, N=65812]

=====

Symbols for Mathematical Operations

A α alpha	I ι iota	P ρ rho
B β beta	K κ kappa	Σ σ sigma
Γ γ gamma	Λ λ lambda	T τ tau
E ε epsilon	M μ mu	Υ υ upsilon
Δ δ delta	N ν nu	Φ φ phi
Z ζ zeta	Ξ ξ xi	X χ chi
H η eta	Ο ο omicron	Ψ ψ psi
Θ θ theta	Π π pi	Ω ω omega